# SEFSC Computation of Uncertainty for Southeast Regional Headboat Survey and Total Recreational Landings Estimates, with Applications to SEDAR 68 Scamp and Yellowmouth Grouper 

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## SEDAR68-DW-31

27 October 2020


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Please cite this document as:

Nuttall, Matthew A., Kyle Dettloff, Kelly E Fitzpatrick, Kenneth Brennan, and Vivian M Matter. 2020. SEFSC Computation of Uncertainty for Southeast Regional Headboat Survey and Total Recreational Landings Estimates, with Applications to SEDAR 68 Scamp and Yellowmouth Grouper. SEDAR68-DW-31. SEDAR, North Charleston, SC. 12 pp.

# SEFSC Computation of Uncertainty for Southeast Region Headboat Survey and Total Recreational Landings Estimates, with Application to SEDAR 68 Scamp and Yellowmouth Grouper 

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October-02-2020

## Introduction

The Southeast Fisheries Science Center (SEFSC) routinely provides stock assessment analysts with estimates of recreational catch and associated measures of uncertainty. The data used to inform these estimates are primarily collected by four surveys in the southeast region (i.e., Texas to North Carolina):

1) Marine Recreational Information Program (MRIP; Matter and Nuttall 2020c)
2) Southeast Region Headboat Survey (SRHS; Fitzpatrick et al. 2017)
3) Texas Marine Sport-Harvest Monitoring Program (TPWD; Nuttall and Matter 2020)
4) Louisiana Creel Survey (LACreel; LDWF 2017, 2020)

Regional estimates for private and rental boats (PR), shore-based anglers (SH), and for-hire guide and charter boat fishing (CH) are largely informed by MRIP survey data (1981+). However, estimates for PR and CH anglers in Texas and PR/SH and CH anglers in Louisiana are, respectively, informed by TPWD (1983+) and the LACreel survey (2014+). Catches of PR, SH , and CH fishing in all other southeastern states and time periods are provided by MRIP: Mississippi to North Carolina (1981+) and Louisiana (1981-2013). Regional estimates for headboats (HB) are primarily estimated from the SRHS, which has operated along the southeast Atlantic since 1972 and in the Gulf of Mexico since 1986. Headboat sampling was initially included in the MRIP survey design, but was discontinued in the southeast in 1986 to avoid overlap with SRHS HB estimates. HB catch is therefore estimated from SRHS data in the South Atlantic (1981+) but from both MRIP (1981-1985) and SRHS data (1986+) in the Gulf of Mexico.

The SEFSC also provides the associated coefficients of variation for MRIP catch estimates, which are commonly applied as measures of uncertainty in regional stock assessments. The
coefficients of variation for MRIP catch totals ( $\widehat{C V}_{M R I P}$ ) are discussed in Dettloff et al. (2020) and summarized below. However, this working paper is the first to introduce the method by which the SEFSC estimates uncertainty for SRHS catch ( $\left.\widehat{C V}_{S R H S}\right)$. Additionally, stock assessment scientists regularly aggregate fishing fleets to reduce model complexity, requiring aggregation of the catch and uncertainty estimates calculated from these two surveys. This working paper also introduces the method by which the SEFSC estimates uncertainty for total recreational catch $\left(\widehat{C V}_{T}\right)$, which is a combination of $\widehat{C V}_{M R I P}$ and $\widehat{C V}_{S R H S}$. The SEFSC does not provide uncertainty measures for TPWD (1983+) and LACreel (2014+) catch estimates. Future efforts may be directed towards applying the $\widehat{C V}_{M R I P}$ approach (Dettloff et al. 2020) to estimate $\widehat{C V}_{T P W D}$ and $\widehat{C V}_{\text {LACreel }}$, but uncertainty in Texas catches and in recent Louisiana catches are currently assumed represented by the $\widehat{C V}_{M R I P}$ estimates from Gulf of Mexico states.

## Methods

## Uncertainty in MRIP Catch

The SEFSC provides annual estimates of catch uncertainty for each fishing mode covered by the MRIP survey (e.g., $\widehat{C V}_{C H}, \widehat{C V}_{P R}$, and $\widehat{C V}_{S H}$ ) and for the total annual catch across all MRIP modes (i.e., $\widehat{C V}_{M R I P}$ ) (Dettloff et al. 2020). These uncertainty measures are calculated using standard survey methodology (Woodruff 1971, Lohr 2010) directly from the raw data to ensure estimates are appropriate for custom (assessment-specific) data aggregations.

For each $g$ aggregation of MRIP catch data, the variance in MRIP catch is calculated as:
Equation (1)

$$
\widehat{V}\left(\hat{C}_{g}\right)=\sum_{h=1}^{n_{g}} \frac{n_{h}}{n_{h}-1} \sum_{i=1}^{n_{h}}\left(y_{h i}-\bar{y}_{h}\right)^{2}
$$

where MRIP catch observations for a given species at the PSU-level (primary sampling unit) $\left(y_{h i}\right)$ are used to calculate stratum level variance estimates, with strata defined by unique combinations of year, mode, wave, and state. The number of observations within each $h$ stratum is represented by $n_{h}$ and the number of strata within each aggregation is represented by $n_{g}$. These estimates are then summed within unique combinations of year and mode to produce annual estimates of uncertainty for each fishing mode (e.g., $\widehat{C V}_{C H}, \widehat{C V}_{P R}$, or $\widehat{C V}_{S H}$ ). Estimates of variance and total catch for each mode are further summed across modes when calculating $\widehat{C V}_{M R I P}$. To avoid low-biased or zero-variance estimates, the SEFSC also substitutes the grand mean of the custom aggregation for $\bar{y}_{h}$ in strata with only a single primary sampling unit (i.e., "lonely" PSUs).

The SEFSC converts these variances into CVs using:
Equation (2)

$$
\widehat{C V}_{g}=\hat{\sigma}_{g} / \hat{C}_{g}
$$

where $\hat{C}_{g}$ is the total estimated catch of a given species in each $g$ aggregation of MRIP data. These mode-specific (e.g., $\widehat{C V}_{C H}, \widehat{C V}_{P R}$, and $\widehat{C V}_{S H}$ ) and total catch $C V s$ (i.e., $\widehat{C V}_{M R I P}$ ) are provided to regional stock assessment scientists as a measure of uncertainty for MRIP catch.

## Uncertainty in SRHS Landings

Industry-reported trip data (i.e., logbooks) are used by the SRHS to estimate the total landings of HB anglers operating throughout the southeast region (Fitzpatrick et al. 2017). This estimation includes a correction factor to account for non-reporting and misreporting of HB fishing activity in the SRHS logbooks, which were originally designed to census all HB fishing activity in the region. Vessel compliance is calculated as the fraction of the total number of fishing trips taken by an individual HB vessel (estimated from the SRHS headboat activity report) for which a corresponding landings record is available from the SRHS logbooks. Correction factors ( $K_{a m v}$ ) are calculated as the inverse of vessel compliance and used to scale vessel-level logbook landings ( $y_{a m v}$ ) into estimates for each $v$ vessel in each SRHS strata, defined by year, area (a), and month $(m)$. Note that total estimated landings ( $\tilde{y}_{a m v}$ ) include both reported logbook data (i.e., $y_{a m v}$, with $K_{a m v} \geq 1$ ) and imputed records for vessels missing catch reports entirely, in which case records are substituted from observations in similar strata (with $K_{a m v}=1$ ). As an example where $g$ identifies landings for a given species at the year level (i.e., $\widehat{C V}_{S R H S}$ ), total headboat landings are estimated as:

Equation (3)

$$
\hat{C}_{g}=\sum_{a=1}^{n} \sum_{m=1}^{n_{a}} \sum_{v=1}^{n_{a m}} \tilde{y}_{a m v} * K_{a m v}
$$

where logbook landings estimates are aggregated across vessels, areas, and months. The total number of headboat fishing vessels within a particular $a-m$ strata is represented by $n_{a m}$, the number of months within a given area by $n_{a}$, and total number of headboat areas by $n$.

As an associated measure of uncertainty for SRHS landings estimates (i.e., $\widehat{C V}_{S R H S}$ ), the SEFSC calculates the variance in reported landings from SRHS logbooks ( $y_{a m v}$ ) across all $v$ vessels within each SRHS strata. These variances are additive and therefore summed across areas and months to estimate the total variance in landings for a given species and year:

Equation (4)

$$
\widehat{V}\left(\hat{C}_{g}\right)=\sum_{a=1}^{n} \sum_{m=1}^{n_{a}}\left(\frac{N_{a m}{ }^{2}}{n_{a m}}\right) \frac{\left(1-n_{a m} / N_{a m}\right)}{n_{a m}-1} \sum_{v=1}^{n_{a m}}\left(\left(y_{a m v} * K_{a m v}\right)-\frac{\sum_{v=1}^{n_{a m}}\left(y_{a m v} * K_{a m v}\right)}{n_{a m}}\right)^{2}
$$

where $N_{a m}$ is the actual number of active fishing vessels within a given $a-m$ strata and $n_{a m}$ the number of active vessels for which the SRHS logbooks contain reported landings data. The $\left(1-n_{a m} / N_{a m}\right)$ term in Equation (4) is a finite population correction factor to account for the
relative coverage of regional headboat fishing effort by the SRHS logbook program and goes to zero when SRHS logbooks are a census of the entire HB fishery. The $\left(N_{a m}{ }^{2} / n_{a m}\right)$ term is a scalar to account for vessels for which catch data is completely missing.

The variance estimation approach described in Equation (4) assumes HB vessels with no reported landings are missing completely at random (MCAR). Note that such an assumption is not made for SRHS landings estimates (Equation 3), within which missing observations of HB landings are imputed from estimated catch records from similar strata, assuming vessels are missing at random (MAR). When converting SRHS variances into CVs (Equation 2), total landings $\left(\hat{C}_{g}\right)$ is estimated from reported landings estimates $\left(y_{a m v}\right)$ under the MCAR assumption for consistency with the variance calculation, rather than using the imputed $\tilde{y}_{a m v}$ totals from Equation (3). However, the imputed estimates of total landings (Equation 3) are provided to analysts for use in stock assessments, not the reported $y_{a m v}$ estimates. Future efforts may be directed towards incorporating a more formal imputation approach (e.g., hot deck) for missing records of SRHS landings to relax the MCAR assumption in Equation (4) and increase consistency between SRHS landings and uncertainty estimates.

## Uncertainty in Total Landings

In assessments where all four recreational fishing modes are combined, total recreational catch $\left(\hat{C}_{T}\right)$ is simply the sum of the individual catch estimates:

Equation (5)

$$
\begin{gathered}
\hat{C}_{T}=\hat{C}_{M R I P}+\hat{C}_{S R H S} \\
\hat{C}_{M R I P}=\hat{C}_{C H}+\hat{C}_{P R}+\hat{C}_{S H} \quad \hat{C}_{S R H S}=\hat{C}_{H B}
\end{gathered}
$$

where $\hat{C}_{M R I P}$ may also contain a $\hat{C}_{H B}$ component if the SRHS does not cover the entire time period (e.g., 1981-1985 for Gulf of Mexico assessments). Similarly, the variances of $\hat{C}_{\text {MRIP }}$ and $\hat{C}_{S R H S}$ are also additive given the estimates are generated from independent surveys for distinct fishing fleets. The variance sum law is then applied to calculate the variance of total recreational catch $\left(\widehat{V}\left(\hat{C}_{T}\right)\right)$ :

Equation (6)

$$
\widehat{V}\left(\hat{C}_{T}\right)=\widehat{V}\left(\hat{C}_{M R I P}\right)+\widehat{V}\left(\hat{C}_{S R H S}\right)
$$

where $\hat{V}\left(\hat{C}_{M R I P}\right)$ is calculated from Equation (1) and $\hat{V}\left(\hat{C}_{S R H S}\right)$ from Equation (4). This approach was taken in SEDAR 49 (2016) to calculate catch uncertainties for various data-limited species caught by a combined commercial and recreational fleet. Substituting $C V s$ for each $\hat{V}\left(\hat{C}_{g}\right)$ term (see Equation 2) and solving for $\widehat{C V}_{T}$, we have:

Equation (7)

$$
\widehat{C V}_{T}=\frac{\sqrt{\left(\widehat{C V}_{M R I P} * \hat{C}_{M R I P}\right)^{2}+\left(\widehat{C V}_{S R H S} * \hat{C}_{S R H S}\right)^{2}}}{\hat{C}_{T}}
$$

which the SEFSC provides as a measure of uncertainty for total recreational catch estimates. Although CVs are only calculated from MRIP and SRHS data, total recreational catch in the Gulf of Mexico (i.e., AB1 in Table 3) includes catch estimates from TPWD and LACreel.

Note that while this section was written with the objective of estimating uncertainty for total recreational catch (i.e., $C V_{T}$ ), it can easily be adapted for custom (assessment-specific) aggregations of fishing modes. For example, considering only charter boat estimates within the MRIP components of Equations (1) and (5) would provide uncertainty estimates for a total forhire fleet, composed of MRIP charterboat and SRHS headboat anglers.

## Results - South Atlantic and Gulf of Mexico Scamp and Yellowmouth Grouper

## SRHS Landings and Uncertainty

Regional estimates of SRHS landings and associated uncertainties are presented for SEDAR 68 scamp and yellowmouth grouper in Figure 1 and Tables 1 and 2. SRHS uncertainty was low overall (grand mean $\widehat{C V}_{S R H S}=0.022$ ), but slightly higher in the Gulf of Mexico ( $\max =0.102$, mean $=0.026$ ) than in the South Atlantic $(\max =0.059$, mean $=0.019)$. This result is largely a function of the census correction factor in Equation (4), which goes to zero in years when the SRHS covers the entire headboat fleet and reported landings can be treated as realized landings. Coverage has generally been high over the entire time series (mean $\approx 80 \%$; Tables 1 and 2 ) and has been close to $100 \%$ since 2010, as represented by the tight error bars over the last decade (Figure 1). This corresponds with an increase in compliance which resulted from enforcement of mandatory reported regulations, which began in 2008 (Fitzpatrick et al. 2017).


Figure 1. Regional estimates of SRHS landings for scamp and yellowmouth grouper with associated uncertainties. Estimates are provided for the South Atlantic (AT) and Gulf of Mexico (GU) with landings calculated from Equation (3) and uncertainties from Equation (4). Error bars represent standard errors of SRHS catch.

Table 1. SRHS landings (numbers of fish) of Gulf of Mexico scamp and yellowmouth grouper with associated uncertainties (CV). Sample size (SS) is provided as the number of active headboat vessels fishing in a given year ( $N_{\text {am }}$; outside parentheses) and those active vessels for which SRHS logbooks contain reported landings data ( $n_{a m}$; inside parentheses). The relative coverage of the headboat fleet by the $\operatorname{SRHS}\left({ }^{\left.n_{a m} / N_{a m}\right)}\right.$ ) is also provided.

| Year | Landings | CV | SS | Coverage |
| :---: | :---: | :---: | :---: | :---: |
| 1986 | 7,263 | 0.058 | $666(425)$ | 0.638 |
| 1987 | 4,577 | 0.046 | $660(463)$ | 0.702 |
| 1988 | 3,399 | 0.049 | $722(568)$ | 0.787 |
| 1989 | 9,310 | 0.030 | $756(584)$ | 0.772 |
| 1990 | 2,388 | 0.033 | $773(714)$ | 0.924 |
| 1991 | 2,056 | 0.019 | $814(669)$ | 0.822 |
| 1992 | 1,611 | 0.011 | $709(703)$ | 0.992 |
| 1993 | 1,685 | 0.006 | $792(773)$ | 0.976 |
| 1994 | 1,137 | 0.040 | $833(765)$ | 0.918 |
| 1995 | 1,370 | 0.102 | $793(646)$ | 0.815 |
| 1996 | 813 | 0.069 | $732(621)$ | 0.848 |
| 1997 | 1,165 | 0.041 | $707(609)$ | 0.861 |
| 1998 | 1,241 | 0.026 | $746(590)$ | 0.791 |
| 1999 | 1,064 | 0.021 | $702(392)$ | 0.558 |
| 2000 | 1,028 | 0.030 | $657(455)$ | 0.693 |
| 2001 | 616 | 0.032 | $662(437)$ | 0.660 |
| 2002 | 705 | 0.046 | $564(445)$ | 0.789 |
| 2003 | 675 | 0.026 | $611(444)$ | 0.727 |
| 2004 | 1,315 | 0.063 | $619(498)$ | 0.805 |
| 2005 | 1,075 | 0.018 | $616(542)$ | 0.880 |
| 2006 | 589 | 0.040 | $598(511)$ | 0.855 |
| 2007 | 668 | 0.042 | $638(500)$ | 0.784 |
| 2008 | 608 | 0.016 | $623(564)$ | 0.905 |
| 2009 | 598 | 0.005 | $659(644)$ | 0.977 |
| 2010 | 992 | 0.005 | $614(590)$ | 0.961 |
| 2011 | 815 | 0.000 | $631(630)$ | 0.998 |
| 2012 | 1,096 | 0.000 | $644(644)$ | 1.000 |
| 2013 | 1,388 | 0.001 | $629(627)$ | 0.997 |
| 2014 | 2,100 | 0.000 | $606(606)$ | 1.000 |
| 2015 | 2,613 | 0.000 | $657(657)$ | 1.000 |
| 2016 | 1,730 | 0.000 | $658(658)$ | 1.000 |
| 2017 | 1,537 | 0.000 | $703(701)$ | 0.997 |
| 2018 | 1,866 | 0.000 | $684(684)$ | 1.000 |
|  |  |  |  |  |
| 18 |  |  |  |  |

Table 2. SRHS landings (numbers of fish) of South Atlantic scamp and yellowmouth grouper with associated uncertainties (CV). Sample size (SS) is provided as the number of active headboat vessels fishing in a given year ( $N_{\text {am }}$; outside parentheses) and those active vessels for which SRHS logbooks contain reported landings data ( $n_{a m}$; inside parentheses). The relative coverage of the headboat fleet by the $\operatorname{SRHS}\left({ }^{\left.n_{a m} / N_{a m}\right)}\right.$ ) is also provided.

| Year | Landings | CV | SS | Coverage |
| :---: | :---: | :---: | :---: | :---: |
| 1981 | 7,286 | 0.059 | $779(513)$ | 0.659 |
| 1982 | 8,021 | 0.025 | $810(502)$ | 0.620 |
| 1983 | 8,510 | 0.027 | $787(523)$ | 0.665 |
| 1984 | 7,317 | 0.020 | $783(499)$ | 0.637 |
| 1985 | 8,472 | 0.050 | $772(506)$ | 0.655 |
| 1986 | 8,241 | 0.026 | $839(608)$ | 0.725 |
| 1987 | 14,868 | 0.018 | $901(641)$ | 0.711 |
| 1988 | 14,128 | 0.020 | $863(546)$ | 0.633 |
| 1989 | 12,098 | 0.058 | $810(452)$ | 0.558 |
| 1990 | 18,095 | 0.022 | $789(414)$ | 0.525 |
| 1991 | 23,650 | 0.034 | $1022(432)$ | 0.423 |
| 1992 | 12,551 | 0.046 | $806(779)$ | 0.967 |
| 1993 | 9,253 | 0.030 | $855(775)$ | 0.906 |
| 1994 | 12,296 | 0.016 | $850(745)$ | 0.876 |
| 1995 | 15,585 | 0.008 | $798(709)$ | 0.888 |
| 1996 | 9,397 | 0.024 | $823(631)$ | 0.767 |
| 1997 | 14,206 | 0.004 | $839(694)$ | 0.827 |
| 1998 | 16,095 | 0.023 | $819(637)$ | 0.778 |
| 1999 | 17,697 | 0.027 | $823(551)$ | 0.670 |
| 2000 | 13,162 | 0.012 | $844(569)$ | 0.674 |
| 2001 | 10,549 | 0.021 | $790(530)$ | 0.671 |
| 2002 | 10,514 | 0.022 | $721(442)$ | 0.613 |
| 2003 | 10,065 | 0.013 | $672(436)$ | 0.649 |
| 2004 | 13,254 | 0.025 | $738(528)$ | 0.715 |
| 2005 | 8,244 | 0.018 | $712(480)$ | 0.674 |
| 2006 | 10,571 | 0.021 | $684(451)$ | 0.659 |
| 2007 | 16,741 | 0.025 | $778(502)$ | 0.645 |
| 2008 | 5,061 | 0.017 | $777(691)$ | 0.889 |
| 2009 | 3,622 | 0.016 | $796(753)$ | 0.946 |
| 2010 | 3,285 | 0.005 | $772(750)$ | 0.972 |
| 2011 | 2,020 | 0.010 | $748(727)$ | 0.972 |
| 2012 | 2,075 | 0.000 | $731(726)$ | 0.993 |
| 2013 | 1,790 | 0.002 | $734(717)$ | 0.977 |
| 2014 | 1,837 | 0.000 | $749(747)$ | 0.997 |
| 2015 | 2,223 | 0.000 | $742(723)$ | 0.974 |
| 2016 | 1,782 | 0.003 | $747(703)$ | 0.941 |
| 2017 | 1,669 | 0.000 | $656(656)$ | 1.000 |
| 2018 | 1,123 | 0.008 | $639(636)$ | 0.995 |
|  |  |  |  |  |

## Total Recreational Landings and Uncertainty

Estimates of total recreational landings and associated uncertainties $\left(\widehat{C V}_{T}\right)$ are presented for SEDAR 68 scamp and yellowmouth grouper in Figure 2 and Tables 3 and 4. Estimates of $\widehat{C V}_{T}$ are large compared to $\widehat{C V}_{S R H S}$, a function of the relatively large estimates of $\widehat{C V}_{M R I P}$ (Table 3 in Matter and Nuttall 2020a, 2020b). The relative contribution of each survey to $\widehat{C V}_{T}$ is a function of their associated catch estimates (Equation 7). Because MRIP-sampled gears landed twenty times more scamp and yellowmouth grouper than SRHS headboats in the Gulf of Mexico (1.9 vs. 42.0 thousand fish annually), $\widehat{C V}_{S R H S}$ contributes little to $\widehat{C V}_{T}$ in this region, which is approximately equal to $\widehat{C V}_{M R I P}$. The relative landings of scamp and yellowmouth grouper are more balanced in the South Atlantic ( 9.4 vs. 16.1 thousand fish), which gives more weight to the annual $\widehat{C V}_{S R H S}$ estimates and results in tighter error bars in the total recreational catch estimates.


Figure 2. Regional estimates of total recreational landings for scamp and yellowmouth grouper with associated uncertainties. Estimates are provided for the South Atlantic (AT) and Gulf of Mexico (GU) with landings calculated from Equation (5) and uncertainties from Equation (6). Error bars represent standard errors of total recreational catch.

Table 3. Total recreational landings (numbers of fish) of Gulf of Mexico scamp and yellowmouth grouper with associated uncertainties (CV).

| Year | AB1 | CV_AB1 |
| :---: | :---: | :---: |
| 1981 | 54,543 | 0.500 |
| 1982 | 102,620 | 0.340 |
| 1983 | 50,387 | 0.520 |
| 1984 | 10,247 | 0.460 |
| 1985 | 21,534 | 0.630 |
| 1986 | 55,038 | 0.276 |
| 1987 | 73,093 | 0.600 |
| 1988 | 42,926 | 0.258 |
| 1989 | 27,921 | 0.247 |
| 1990 | 8,912 | 0.650 |
| 1991 | 16,929 | 0.633 |
| 1992 | 15,262 | 0.331 |
| 1993 | 25,119 | 0.401 |
| 1994 | 14,005 | 0.533 |
| 1995 | 5,700 | 0.486 |
| 1996 | 13,126 | 0.553 |
| 1997 | 15,885 | 0.334 |
| 1998 | 21,974 | 0.198 |
| 1999 | 40,794 | 0.273 |
| 2000 | 11,591 | 0.336 |
| 2001 | 14,190 | 0.201 |
| 2002 | 25,168 | 0.272 |
| 2003 | 46,067 | 0.503 |
| 2004 | 53,423 | 0.254 |
| 2005 | 62,358 | 0.472 |
| 2006 | 105,979 | 0.766 |
| 2007 | 41,129 | 0.305 |
| 2008 | 60,457 | 0.455 |
| 2009 | 49,845 | 0.553 |
| 2010 | 28,399 | 0.454 |
| 2011 | 44,764 | 0.255 |
| 2012 | 77,288 | 0.335 |
| 2013 | 78,538 | 0.246 |
| 2014 | 78,436 | 0.282 |
| 2015 | 108,607 | 0.517 |
| 2016 | 70,282 | 0.322 |
| 2017 | 47,987 | 0.406 |
| 2018 | 56,105 | 0.328 |
|  |  |  |

Table 4. Total recreational landings (numbers of fish) of South Atlantic scamp and yellowmouth grouper with associated uncertainties (CV).

| Year | AB1 | CV_AB1 |
| :---: | :---: | :---: |
| 1981 | 21,332 | 0.586 |
| 1982 | 18,472 | 0.407 |
| 1983 | 9,558 | 0.071 |
| 1984 | 17,971 | 0.285 |
| 1985 | 14,767 | 0.346 |
| 1986 | 11,145 | 0.152 |
| 1987 | 16,395 | 0.048 |
| 1988 | 33,179 | 0.207 |
| 1989 | 31,389 | 0.210 |
| 1990 | 44,079 | 0.230 |
| 1991 | 34,041 | 0.115 |
| 1992 | 27,071 | 0.200 |
| 1993 | 28,647 | 0.244 |
| 1994 | 45,068 | 0.218 |
| 1995 | 15,680 | 0.010 |
| 1996 | 16,991 | 0.255 |
| 1997 | 17,370 | 0.149 |
| 1998 | 19,820 | 0.065 |
| 1999 | 25,625 | 0.116 |
| 2000 | 42,917 | 0.257 |
| 2001 | 25,255 | 0.181 |
| 2002 | 58,482 | 0.213 |
| 2003 | 45,269 | 0.296 |
| 2004 | 40,968 | 0.264 |
| 2005 | 34,926 | 0.512 |
| 2006 | 52,551 | 0.407 |
| 2007 | 59,190 | 0.222 |
| 2008 | 31,888 | 0.286 |
| 2009 | 18,105 | 0.400 |
| 2010 | 11,148 | 0.317 |
| 2011 | 6,867 | 0.353 |
| 2012 | 9,073 | 0.363 |
| 2013 | 10,584 | 0.341 |
| 2014 | 40,226 | 0.897 |
| 2015 | 7,453 | 0.421 |
| 2016 | 8,590 | 0.396 |
| 2017 | 13,976 | 0.740 |
| 2018 | 4,068 | 0.347 |
|  |  |  |

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