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# Efficacy of TIP length composition for use in length-based mortality estimation 

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## Executive summary

We examine statistical properties of TIP length composition data for spiny lobster including the spatial distribution of sampling effort relative to fishing activity and the apparent reliability of mean length estimates based on sample size. Our analysis evaluates the information content in TIP length compositions, prior to the use of these data in stock assessment. Correspondence between fishery activity (catches and trips taken) and sampling of those fishing events via TIP was reasonable for Puerto Rico, St. Thomas, and St. Croix. Analysis of effective sample sizes needed to estimate mean length from TIP datasets were reasonable for Puerto Rico and St. Croix, while ESS for St. Thomas was more concerning. Simulation modeling also reasonably supported the reliability of TIP length composition datasets for use in data-limited assessment methods.

## Introduction

Some data-limited stock assessment methods rely on the observation of mean length to estimate total mortality and related measures of a stock's reproductive state (Beverton and Holt, 1957; Ehrhardt and Ault, 1992; Gedamke and Hoenig, 2006). The reliability of mean length observations will influence the corresponding reliability of total mortality estimates. Accordingly, we examined the statistical properties of TIP length composition data for spiny lobster prior to considering whether to use these data in stock assessment. We approached the question of length composition efficacy for data-limited stock assessment by first graphically comparing fishery catches to TIP sampling intensity, by both gear and spatial region. We then evaluated the effective sample size (ESS) of TIP data, both spatially and temporally. ESS reveals the extent to which the observed sample size, $n$, is consistent with a random sample of the statistical population, ESS. Thus, ESS is a measure of the information content in sample of length observations. Finally, we conducted simulation of length composition data according to specified ESSs to reveal the correspondence between ESS and accuracy of mean length observations.

## Graphical representation of TIP sampling by gear and coast

## Methods

We examined the gears associated with spiny lobster sampled by the TIP datasets for Puerto Rico, St. Thomas, and St. Croix. Percentages of samples by island and gear types in the TIP data were compared to the percentages of pounds reported by island and gear types in the logbook data (Figure 1). For Puerto Rico, we also looked at the percentages by gear type and coast (Figure 2). Because of the two different units used to develop annual percentages (pounds from the logbook data and individuals sampled by TIP), caution when interpreting these plots is warranted. However, they can be used to get a preliminary sense of the representativeness of the TIP data for spiny lobster.

## Results

In Puerto Rico and in St. Thomas, the predominant gears associated with the majority of spiny lobster landings were also associated with the majority of spiny lobsters sampled by TIP. In Puerto Rico, some minor differences exist between the two data sources. Compared to the percentages of landings, TIP samples from diving are slightly underrepresented and the TIP samples from pots and traps are slightly overrepresented. However, these differences are not of particular concern for using the length data.

In St. Croix, a disproportionately large percentage of spiny lobster TIP samples are associated with the pot and trap gear type compared to the percentages of landings for the same gear. However, this difference is not of particular concern since the analysis will focus on the predominant gear (Diving).

## Effective sample size

## Methods

We examined the precision of length-frequency distributions of spiny lobster obtained from TIP datasets for Puerto Rico, St. Thomas, and St. Croix. Given complications of field sampling, including those related to the use of fishery-dependent sampling, length samples collected from $n$ sampling events are not necessarily a random sample from the entire population, but instead may be subject to errors attributable to data collection methods (Hulson et al., 2012). This circumstance can lead to less information about the population being contained in the $M$ total length observations than would have been obtained from sampling $M$ fish randomly from the population (Pennington et al., 2002). The quantity known as effective sample size (ESS) is an estimate of the sample size obtained through random sampling that would contain the equivalent information contained in the observed sample, $M$. The ESS is less than or equal to $M$. Steps necessary to estimate ESS are found in Pennington et al. (2002), where in our application to the TIP datasets, each fishing trip was treated as sampling event, $n$. ESS was calculated separately for each of the three island platforms and for each year. For Puerto Rico, we also estimated ESS for four spatial areas.

Simulation was carried out to evaluate the effect of ESS on precision and bias of mean length estimates arising from sampling of length composition datasets. Length composition datasets were generated from a population simulation representative of spiny lobster. The simulated population was used to generate equilibrium length composition, binned using 2 mm intervals. Length-structured population dynamics were simulated, which is an approach well-suited for modeling marine invertebrates (Haddon, 2011). Length-based models account for survival, growth, and reproduction through time by assigning individuals to length classes or length bins. Numbers-at-length matrices differ from numbers-at-age matrices because the latter tracks specific cohorts as they transition between age classes, while the former probabilistically tracks transitions between length classes where individuals from several cohorts are likely to be found in any given length bin (Haddon, 2011). The simulated spiny lobster population had life history characteristics representative of US Caribbean stocks: segmented growth obtained through analysis of mark-recapture data from St. Thomas, natural mortality of 0.34 year $^{-1}$ (FAO, 2001), maturity-at-length obtained from TIP data (Die, 2005), length-weight from TIP data (SEDAR, 2018), and Beverton-Holt stock-recruitment relationship with assumed steepness of 0.95 . This population simulation was used to generate binned equilibrium length composition at depletion levels of 0.8 and 0.2 .

Given simulated 'true' length compositions, a sample of these lengths was drawn from numbers-at-lengths above the minimum harvest size of 89 mm using a multinomial sampling distribution, with a specified sample size (which is equivalent to ESS). The mean of the length sample was calculated, along with standard deviation of mean estimator (i.e., $S D_{i} / \sqrt{E S S_{i}}$, where SD is standard deviation of the sample, and ESS is number of length samples). Percent bias between the estimated mean and true mean of the simulated length composition was calculated: (estimated mean - true mean) / true mean x 100 . This process was repeated for 1,000 independent replicates at a given ESS. We then repeated this sampling exercise for ESSs ranging
from 10 to 500. Results were summarized according to mean and centered $95 \%$ distributions in bias in mean length versus ESS. We similarly summarized the precision of the mean estimator as coefficient of variation versus ESS.

## Results

Effective sample sizes (ESS) for all of the island platforms vary considerably through time (Tables $1 \& 2$ ). For Puerto Rico, East, West, and South regions tended to have substantially higher ESS than the North region, as the former three regions represent the vast majority of the area where fishing takes place. By pooling length samples across all four regions, each year had ESS $>50$, with the exception of 1988. Temporal trends in ESS were similarly stable for St. Croix, with all years having ESS > 50 with the exception of 2011. St. Thomas length sampling had similar consistency. Note that for each island platform, missing values reflect a lack of data availability in those respective years.

Simulations highlight that bias in estimation of mean length can exceed $10 \%$ at very low ESS (Fig. 3). However, exponential reduction in bias occurs as ESS increases, and bias tends to be less than 5\% at ESSs exceeding 50. Similarly, imprecision of the mean estimator is, on average, $2 \%$ or less when ESS exceeds 50. These results are encouraging, given the observed ESS estimates from TIP sampling for Puerto Rico and St. Croix.

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Table 1. Effective sample size (ESS) for St. Croix (STX) and St. Thomas (STT) TIP length composition datasets. n is number of length observations.

| Year | STX |  | STT |  |
| :---: | :---: | :---: | :---: | :---: |
|  | n | ESS | n | ESS |
| 1980 |  |  | 301 | 61 |
| 1981 | 260 | 62 | 916 | 311 |
| 1982 | 408 | 201 |  |  |
| 1983 | 399 | 247 | 103 | 49 |
| 1984 | 867 | 237 | 755 | 102 |
| 1985 | 1631 | 991 | 1114 | 294 |
| 1986 | 1329 | 1017 | 299 | 111 |
| 1987 | 1621 | 713 | 415 | 71 |
| 1988 | 761 | 695 | 701 | 492 |
| 1989 | 54 | 135 |  |  |
| 1990 | 552 | 407 |  |  |
| 1991 | 625 | 223 |  |  |
| 1992 | 798 | 297 | 193 | 51 |
| 1993 | 824 | 487 | 212 | 142 |
| 1994 | 790 | 441 | 76 | 226 |
| 1995 | 547 | 154 | 29 | 6 |
| 1996 | 400 | 258 | 117 | 291 |
| 1997 | 634 | 245 |  |  |
| 1998 | 517 | 257 |  |  |
| 1999 | 542 | 257 |  |  |
| 2000 | 320 | 61 |  |  |
| 2001 | 295 | 212 |  |  |
| 2002 | 598 | 252 | 318 | 60 |
| 2003 | 564 | 269 | 354 | 48 |
| 2004 | 433 | 188 | 162 | 65 |
| 2005 | 526 | 435 | 203 | 181 |
| 2006 | 366 | 330 | 799 | 179 |
| 2007 | 647 | 548 |  |  |
| 2008 | 404 | 124 | 90 | 21 |
| 2009 | 388 | 231 | 738 | 114 |
| 2010 | 717 | 411 | 933 | 235 |
| 2011 | 33 | 17 | 619 | 252 |
| 2012 |  |  | 444 | 659 |
| 2013 |  |  | 286 | 150 |
| 2014 |  |  |  |  |
| 2015 |  |  | 110 | 8 |
| 2016 | 169 | 77 | 722 | 262 |
| 2017 | 273 | 140 | 619 | 207 |
|  |  |  |  |  |
|  |  |  |  |  |

Table 2. Effective sample size (ESS) for Puerto Rico TIP length composition datasets. n is number of length observations.

| Year | East |  |  | West |  | North |  | South |  | All |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | n | ESS | n | ESS | n | ESS | n | ESS | n | ESS |  |
| 1980 | 131 | 108 | 126 | 55 | 11 | 17 | 59 | 64 | 337 | 170 |  |
| 1981 |  |  |  |  |  |  |  |  |  |  |  |
| 1982 |  |  |  |  |  |  |  |  |  |  |  |
| 1983 | 11 | 6 | 114 | 62 |  |  | 8 | 12 | 133 | 79 |  |
| 1984 | 304 | 237 | 642 | 392 |  |  | 104 | 130 | 1272 | 840 |  |
| 1985 | 148 | 53 | 298 | 235 | 5 | NA | 491 | 420 | 970 | 475 |  |
| 1986 | 84 | 36 | 157 | 57 | 13 | 3 | 221 | 290 | 551 | 248 |  |
| 1987 | 75 | 35 | 104 | 62 |  |  | 313 | 261 | 603 | 424 |  |
| 1988 |  |  |  |  |  |  | 4 | NA | 38 | 18 |  |
| 1989 | 62 | 65 | 51 | 14 | 26 | 23 | 393 | 301 | 545 | 220 |  |
| 1990 | 14 | 17 | 25 | 1 | 33 | 12 | 477 | 457 | 549 | 109 |  |
| 1991 | 185 | 47 | 24 | 25 | 25 | 19 | 691 | 296 | 925 | 344 |  |
| 1992 | 88 | 43 | 175 | 107 | 3 | 3 | 495 | 341 | 779 | 540 |  |
| 1993 | 91 | 25 | 71 | 32 | 8 | 2 | 393 | 186 | 563 | 234 |  |
| 1994 | 62 | 27 | 132 | 93 |  |  | 19 | 12 | 217 | 125 |  |
| 1995 | 92 | 54 | 131 | 60 | 1 | NA | 300 | 198 | 624 | 158 |  |
| 1996 | 204 | 84 | 58 | 36 | 1 | NA | 249 | 176 | 512 | 280 |  |
| 1997 | 13 | 164 | 20 | 8 |  |  | 197 | 122 | 230 | 147 |  |
| 1998 | 258 | 181 | 192 | 115 |  |  | 188 | 87 | 638 | 344 |  |
| 1999 | 314 | 118 | 291 | 44 | 50 | 9 | 366 | 198 | 1021 | 252 |  |
| 2000 | 212 | 62 | 234 | 134 | 12 | 38 | 328 | 213 | 795 | 283 |  |
| 2001 | 506 | 237 | 308 | 226 | 72 | 47 | 392 | 363 | 1278 | 755 |  |
| 2002 | 81 | 23 | 227 | 100 | 46 | 16 | 239 | 123 | 606 | 181 |  |
| 2003 | 687 | 377 | 373 | 199 | 131 | 39 | 256 | 200 | 1447 | 758 |  |
| 2004 | 598 | 196 | 760 | 438 | 104 | 16 | 159 | 132 | 1621 | 416 |  |
| 2005 | 169 | 89 | 974 | 558 | 29 | 11 | 260 | 161 | 1432 | 775 |  |
| 2006 | 181 | 18 | 1041 | 132 | 143 | 62 | 416 | 163 | 1781 | 248 |  |
| 2007 | 262 | 100 | 1014 | 298 | 50 | 36 | 384 | 192 | 1710 | 507 |  |
| 2008 | 131 | 11 | 488 | 309 | 57 | 12 | 201 | 78 | 877 | 173 |  |
| 2009 | 185 | 68 | 873 | 449 | 72 | 50 | 456 | 187 | 1586 | 673 |  |
| 2010 | 147 | 82 | 770 | 386 | 21 | 319 | 399 | 232 | 1337 | 705 |  |
| 2011 | 221 | 104 | 1704 | 985 | 173 | 128 | 495 | 365 | 2593 | 1344 |  |
| 2012 | 318 | 169 | 1867 | 1120 | 299 | 204 | 492 | 296 | 2976 | 1442 |  |
| 2013 | 713 | 244 | 2123 | 1159 | 86 | 28 | 333 | 262 | 3255 | 1537 |  |
| 2014 | 1079 | 682 | 1549 | 870 | 2 | 2 | 972 | 397 | 3615 | 1808 |  |
| 2015 | 854 | 523 | 1349 | 845 | 25 | 5 | 782 | 602 | 3010 | 1749 |  |
| 2016 | 1030 | 296 | 1631 | 931 |  |  | 753 | 448 | 3435 | 1481 |  |
|  |  |  |  |  |  |  |  |  |  |  |  |



Figure 1. Boxplots of annual percentages of pounds landed (red) and of individuals sampled (blue) of spiny lobster by island and gear type.


Figure 2. Boxplots of annual percentages of pounds landed (red) and of individuals sampled (blue) of spiny lobster in Puerto Rico by coast and gear type.


Figure 3. Simulated effects of effective sample size (ESS) on corresponding bias in estimating mean length (A \& C) and precision of the estimated mean (B \& D). Simulations consisted of generated equilibrium length composition at depletion levels of $0.8(\mathrm{~A} \& \mathrm{~B})$ and $0.2(\mathrm{C} \& \mathrm{D})$ for a population with life history characteristics representative of spiny lobster. Length observations were sampled according to a multinomial sampling distribution with ESSs ranging from 10 to 500 (x-axes in all plots).

