# Estimating total mortality rates and calculating overfishing limits from length observations for six U.S. Caribbean stocks 

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# Estimating total mortality rates and calculating overfishing limits from length observations for six U.S. Caribbean stocks 

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## 1. Introduction

The mean length-based estimator of Gedamke and Hoenig (2006) is a non-equilibrium extension of the Beverton-Holt mean length mortality estimator used to estimate total instantaneous mortality rates $Z$. Gedamke and Hoenig (2006) derived the transitional behavior of the population mean length following a change in $Z$ and then generalized the derivation to include multiple changes in total mortality. From a time series of mean length data, total mortality rates are estimated in blocks of time as well as the years in which the mortality changed. The model uses a likelihood approach to obtain parameters that maximize goodness of fit to the mean length data. The mean length estimator assumes:

1. Recruitment is constant over time.
2. Growth is deterministic following a von Bertalanffy growth equation and is timeinvariant.
3. Selectivity is knife-edge above the length of full selectivity $L_{c}$ and is time-invariant.

The mean length estimator requires an external estimate of von Bertalanffy growth parameters $L_{\infty}$ and $K$ and a time series of length measurements for the stock. First, the length of full selectivity $L_{c}$ is obtained from the data. Typically, the $L_{c}$ is selected to be the peak (mode) of the length frequency histogram of data combined for all years in the time series and the annual mean lengths of animals larger than length $L_{c}$ is calculated. Annual length frequency histograms should also be examined to explore trends in the mode of the histogram over time, which would coincide with a change in selectivity or recruitment. A sensitivity analysis should also be performed in which several values of $L_{c}$ are chosen and the mean lengths re-calculated for each alternative value of $L_{c}$. This allows for a test of the constancy of mortality rate over the size range assumed to be fully recruited. For increasing values of $L_{c}$, an increasing trend in the estimates of mortality may suggest dome-shaped selectivity, while no trend in the estimates would indicate that selectivity is flat-topped.

To estimate historical mortality rates, the number of changes in mortality is first specified by the user. The model is fitted multiple times with increasing complexity (more time blocks) until the increase in goodness of fit is no longer statistically significant with increasing complexity or until trends in residuals are removed. In practice, AIC is used to select the best model considering both goodness of fit and model parsimony (Burnham and Anderson 2002).

For a single stock, length data from several fleets or gears may be available, in which case the data from the best-sampled or most abundant gear should be used to prevent conflation of differing selectivity patterns with stock status.

The estimates of total mortality from the mean length estimator can be used to calculate an overfishing limit (OFL) (Bryan et al. in progress). The most recent fishing mortality rate $F_{\text {recent }}$ in the most recent years of the time series is derived as the difference of the total mortality and an external estimate of the natural mortality rate $M$. This benchmark fishing mortality rate can be compared to a reference point $F_{\text {reference }}$. Based on the ratio of the $F_{\text {reference }}$ and $F_{\text {recent }}$, the management procedure adjusts recent catch $C_{\text {recent }}$ concurrent with $F_{\text {recent }}$ to derive an overfishing limit:

$$
\begin{equation*}
\text { OFL }=\frac{F_{\text {reference }}}{F_{\text {recent }}} C_{\text {recent }} . \tag{1}
\end{equation*}
$$

High instantaneous fishing mortality rates should be converted to annual exploitation rates (as a ratio of the total population fished) to calculate the OFL:

$$
\begin{equation*}
u=1-\exp (-F) \tag{2}
\end{equation*}
$$

## 2. Application to U.S. Caribbean species

### 2.1. Mean length estimator

The mean length estimator was used to estimate total mortality rates for six U.S. Caribbean stocks: yellowtail snapper and hogfish in Puerto Rico (PR), queen triggerfish and spiny lobster in St. Thomas/St. John (STT), and spiny lobster and stoplight parrotfish in St. Croix (STX). Length observations were obtained from the Trip Interview Program (TIP) of the Southeast Fisheries Science Center. A summary of the data and gear aggregations is available in Bryan (2015). Life history parameters for the six stocks were obtained from Adams (2015). The modal length from length histograms of all observations was used to select $L_{c}$ for the mean length estimator (Figure 1).

The mean lengths of animals above length $L_{c}$ were calculated, which necessitated discarding of data of lengths below $L_{c}$. Sample sizes for the data used to calculate mean lengths are reported in Table 1. Annual length frequency figures are available in Bryan (2015). From these annual length frequency plots, the annual modal length was also calculated to detect possible changes in selectivity over time (Figure 2). Annual length frequency plots for St . Thomas/St. John spiny lobster from the trap gear and St. Croix spiny lobster from the dive/by hand/spear gear are presented in Figure 3 and Figure 4, respectively. Total mortality rates were estimated assuming up to 3 changes in mortality occurred in the time series. AIC was used to select the best fit model, with the more complex model (more changes in mortality) selected if
the reduction in AIC was greater than 4 units. Sensitivity runs using alternate choices for $L_{c}$ were also used (Tables 2-7). Annual mean lengths $>L_{c}$ were re-calculated and the model fit and selection was repeated for each sensitivity run.

### 2.1.1. Puerto Rico yellowtail snapper

A total of 76,574 observations of yellowtail snapper were available from the handline gear in Puerto Rico from 1983-2013. From all records in this time period, the modal length was 280 mm . However, there appeared to be an increase in the modal length after the year 2000 suggesting a change in selectivity around this time (Figure 2a).

Since there appeared to be no trend in the modal length after 2000, the $L_{c}$ was selected from the modal length from observations sampled after 2000. Thus, estimated mortality rates from the mean length estimator are assumed to be conditional on the more recent selectivity pattern. Using this criterion, the modal length was still 280 mm , which was chosen as the $L_{c}$. Total mortality was estimated to be 0.56 during 1983-1992, decreased to 0.26 during 19932004, and subsequently increased to 0.48 since 2004 (Figure 5). The estimates did not appear to be very sensitive to $L_{c}$, although there was a noticeable decrease most recent estimate of $Z$ when a very high $L_{c}$ was chosen (Table 2).

### 2.1.2. Puerto Rico hogfish

A total of 4,163 observations of hogfish were available from the diving/by hand/spear gear in Puerto Rico from 1983-2013. From all records in this time period, the modal length was 280 mm , which was chosen as the $L_{c}$. Total mortality was estimated to be 0.36 during 1983 1992, increased to 0.56 during 1993-2000, and subsequently decreased to 0.34 since 2001 (Figure 6). The estimates of $Z$ were not particularly sensitive to the choice of $L_{c}$ (Table 3).

### 2.1.3. St. Thomas/St. John queen triggerfish

A total of 8,051 observations of queen triggerfish were available from the trap fishery in St. Thomas/St. John from 1983-2013. The modal length was 320 mm , which was chosen as the $L_{c}$. The mean length estimator suggested equilibrium mortality during 1983-2013, i.e. no changes in mortality (Figure 7). Total mortality was estimated to be 1.34 during this period, although there is a trend in the estimated mortality depending on the choice of $L_{c}$, with higher mortality estimated when a larger $L_{c}$ is used (Table 4).

### 2.1.4. St. Thomas/St. John spiny lobster

A total of 10,384 observations of spiny lobster were available from the trap fishery in St. Thomas/St. John from 1980-2013 (Figure 3). There were no length observations for the years 1982, 1989-1991, 1997-2001, and 2007. Sample sizes were high prior to 1989, then declined before increasing after 2002. The modal length was 100 mm , which was chosen as the $L_{c}$. The
mean length estimator suggested one change in mortality. Total mortality was estimated to be 0.59 prior to 1986 and increased to 0.98 afterwards (Figure 8). There is a slight trend in the estimated mortality depending on the choice of $L_{c}$, although the estimates are similar when the chosen $L_{c}$ is 100 mm and 110 mm (Table 5).

### 2.1.5. St. Croix spiny lobster

A total of 6,674 observations of spiny lobster were available from the diving/by hand/spear fishery in St. Croix from 1981-2013. There are no data for 1986, 1988-1990, and 2011-2013 (Figure 4). Sample sizes increased in the 1990s and 2000s, with a declining trend since 2008. The modal length was 90 mm , which was chosen as the $L_{c}$. The mean length estimator suggested that one increase in total mortality occurred and there appears to be a decrease in the mean length over time. Total mortality was estimated to be 0.35 prior to 1979 and increased to 1.11 afterwards (Figure 9). The model estimated a mortality rate that occurred prior to 1980 to explain the decreasing trend in mean length in the initial years of the time series. Since there are no data prior to 1980, there is high uncertainty in this estimate. However, the relative stability in the mean length after the mid-1980s provided stability in the model to estimate the most recent total mortality rate. There is a trend in the estimated mortality depending on the choice of $L_{c}$, with higher mortality estimated when a larger $L_{c}$ is used (Table 6).

### 2.1.6. St. Croix stoplight parrotfish

A total of 2,605 observations of stoplight parrotfish were available from the diving/by hand/spear fishery in St. Croix. Only 10 intermittent years of length observations were available starting in 1996. The modal length was 270 mm , which was chosen as the $L_{c}$. There was not enough information to detect a change in total mortality. The equilibrium total mortality form 1996 - 2013 was estimated to be 2.31 during this period (Figure 10). There is a trend in the estimated mortality depending on the choice of $L_{c}$, with higher mortality estimated when a larger $L_{c}$ is used (Table 7).

### 2.2. Calculating OFLs

OFLs were calculated for the six stocks using Equation 1 using two reference points, $F_{0.1}$ and $F_{S P R \%}$, as proxies for $F_{\text {reference. }}$. From a yield-per-recruit (YPR) analysis, $F_{0.1}$ is the fishing mortality rate at which the slope of the YPR curve is $10 \%$ of that at the origin. From a spawning-potential-ratio (SPR) analysis, $F_{S P R \%}$ is the fishing mortality rate that reduces the spawning biomass per recruit to a pre-specified percentage of that from unfished conditions. $F_{0.1}$ is generally used if there are concerns with growth overfishing while $F_{S P R \%}$ is used to consider recruitment overfishing (SEDAR 2014). For both, weight-at-age information for the stock is required, while SPR also requires maturity information. The YPR and SPR formulation from SEDAR (2014) was used for the analyses for this assessment. For the YPR and SPR analyses, knife-edge selectivity above length $L_{c}$. Maturity was also assumed to be knife-edge for the SPR


The benchmark fishing mortality rates $F_{\text {recent }}$ were derived as the difference of the most recent total mortality rates from the mean length estimator and external estimates of the natural mortality rate $M$. For the four St. Thomas/St. John and St. Croix stocks, exploitation rates $u$ were used in lieu of high instantaneous fishing mortality rates. Recent catch was calculated as the mean of total catch corresponding to the time interval of the most recent mortality rate estimated by the mean length estimator. For example, for Puerto Rico yellowtail snapper, the change to the latest mortality rate was estimated to occur at 2004.5. Thus, mean catch since 2005 was used as $C_{\text {recent }}$ in Equation 1.

For all stocks, $F_{0.1}$ was lower than $F_{30 \%}$ (Table 8). For St. Thomas/St. John spiny lobster, the spawning potential ratio did not reach $30 \%$ from that at unfished conditions for any practical fishing mortality rate. Thus, a spawning potential ratio of $40 \%$ was considered for this stock. For all stocks, the benchmark $F_{\text {recent }}$ was above the reference $F_{0.1}$ and the respective $O F L_{F 0.1}$ were adjusted downward from the mean catch. Using $F_{S P R \%}$, the $O F L_{F S P R \%}$ was a much smaller reduction from mean catch compared to $O F L_{F 0.1}$ for Puerto Rico hogfish, St. Thomas/St. John queen triggerfish, and St. Croix spiny lobster. For Puerto Rico yellowtail snapper and St. Thomas/St. John spiny lobster, the $O F L_{F S P R \%}$ control rule increased the overfishing limit above the mean catch. For St. Croix stoplight parrotfish, both control rules significantly reduced the mean catch to obtain the OFL.

In addition to the point estimates, a distribution of OFLs was calculated incorporating uncertainty in life history parameters and catch data with the DLMtool package (Carruthers et al. 2014). 500 simulations were performed using $F_{0.1}$ and $F_{S P R \%}\left(F_{40 \%}\right.$ for St. Thomas/St. John spiny lobster, $F_{30 \%}$ for all others) as reference points using assuming lognormal error in natural mortality, growth, maturity, and catch. OFLs were positively skewed, with those obtained from $F_{S P R \%}$ as the reference point larger than those from $F_{0.1}$ (Figure 11). The point estimates for the OFLs were generally close to the median, but for St. Croix stoplight parrotfish using both $F_{0.1}$ and $F_{30 \%}$, where both point estimates were significantly larger than the median.

## 3. Discussion

In general, $F_{S P R \%}$ produced OFLs with smaller deviations from mean catch compared to $F_{0.1}$. The benchmark mortality estimates for Puerto Rico yellowtail snapper and hogfish from the mean length estimator appear consistent with the stocks' respective life histories. These estimates were insensitive across the range of values for $L_{c}$ for the mean length estimator. Using $F_{0.1}$, a large reduction in the mean catch was indicated for the OFL for both stocks. Using $F_{30 \%}$, the reduction was slight for hogfish while there was a large increase of the mean catch for the OFL for yellowtail snapper.

For St. Thomas/St. John and St. Croix spiny lobster, the estimates of total mortality are somewhat high compared to the stocks' life histories. Using $F_{0.1}$, a large reduction in the mean
catch was indicated for the OFL for both stocks. Using $F_{30 \%}$, the reduction was still substantial for St. Croix. Due to the larger full selectivity length $L_{c}$ for the St. Thomas/St. John stock (100 mm vs. 90 mm ), $F_{30 \%}$ could not be obtained. $F_{40 \%}$ was used instead as the SPR proxy, which still provided for a larger OFL relative to mean catch.

For St. Thomas/St. John queen triggerfish and St. Croix stoplight parrotfish, the estimates of total mortality were very high. For queen triggerfish, a large reduction in the mean catch was indicated for the OFL using $F_{0.1}$ while the reduction was more modest using $F_{30 \%}$. St. Croix stoplight parrotfish was a noticeable outlier for the high estimates of total mortality and deviations in the point estimate and the distribution of OFL. In addition to high exploitation, the high estimated total mortality rate could also arise from a positive-biased estimate of the von Bertalanffy $L_{\infty}$ and/or severe dome-shaped selectivity. Both point estimate OFLs indicated a large reduction in catch, but there is more uncertainty as these estimates were not close to the median from the simulations.

## 4. Acknowledgements

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## 6. Tables

Table 1. Annual sample sizes of length observations greater than length $L_{c}$ used in mean length estimator.

| Year | Yellowtail snapper (PR) $L_{c}=280 \mathrm{~mm}$ | $\begin{aligned} & \operatorname{Hogfish}(\mathrm{PR}) \\ & L_{c}=280 \mathrm{~mm} \\ & \hline \end{aligned}$ | $\begin{array}{r} \text { Queen trigger- } \\ \text { fish (STT) } \\ L_{c}=320 \mathrm{~mm} \\ \hline \end{array}$ | $\begin{array}{r} \text { Spiny lobster } \\ (\mathrm{STT}) \\ L_{c}=100 \mathrm{~mm} \\ \hline \end{array}$ | $\begin{array}{r} \text { Spiny lobster } \\ (\mathrm{STX}) \\ L_{c}=90 \mathrm{~mm} \\ \hline \end{array}$ | $\begin{array}{r} \text { Stoplight parrot- } \\ \text { fish (STX) } \\ L_{c}=270 \mathrm{~mm} \\ \hline \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1980 | 0 | 0 | 0 | 164 | 0 | 0 |
| 1981 | 0 | 0 | 0 | 456 | 135 | 0 |
| 1982 | 0 | 0 | 0 | 0 | 72 | 0 |
| 1983 | 19 | 21 | 5 | 43 | 88 | 0 |
| 1984 | 376 | 44 | 332 | 619 | 126 | 0 |
| 1985 | 658 | 60 | 299 | 646 | 6 | 0 |
| 1986 | 225 | 16 | 187 | 239 | 0 | 0 |
| 1987 | 34 | 50 | 50 | 356 | 9 | 0 |
| 1988 | 328 | 82 | 0 | 530 | 0 | 0 |
| 1989 | 141 | 153 | 0 | 0 | 0 | 0 |
| 1990 | 562 | 148 | 0 | 0 | 0 | 0 |
| 1991 | 4420 | 149 | 0 | 0 | 242 | 0 |
| 1992 | 3601 | 108 | 0 | 156 | 438 | 0 |
| 1993 | 2370 | 68 | 111 | 137 | 326 | 0 |
| 1994 | 1695 | 15 | 63 | 53 | 283 | 0 |
| 1995 | 1699 | 43 | 43 | 29 | 92 | 0 |
| 1996 | 183 | 30 | 13 | 87 | 29 | 4 |
| 1997 | 256 | 38 | 0 | 0 | 219 | 0 |
| 1998 | 1132 | 61 | 0 | 0 | 283 | 64 |
| 1999 | 2023 | 135 | 0 | 0 | 272 | 0 |
| 2000 | 2498 | 158 | 0 | 0 | 201 | 0 |
| 2001 | 1285 | 237 | 0 | 0 | 295 | 0 |
| 2002 | 2341 | 103 | 133 | 281 | 570 | 56 |
| 2003 | 3156 | 112 | 43 | 337 | 557 | 60 |
| 2004 | 1571 | 87 | 15 | 117 | 412 | 37 |
| 2005 | 1647 | 93 | 86 | 128 | 435 | 0 |
| 2006 | 1962 | 103 | 109 | 566 | 360 | 171 |
| 2007 | 2441 | 64 | 0 | 0 | 514 | 338 |
| 2008 | 1633 | 45 | 11 | 72 | 158 | 363 |
| 2009 | 2013 | 244 | 348 | 585 | 75 | 488 |
| 2010 | 1699 | 205 | 727 | 768 | 109 | 1 |
| 2011 | 520 | 297 | 431 | 494 | 0 | 0 |
| 2012 | 668 | 222 | 81 | 369 | 0 | 0 |
| 2013 | 266 | 116 | 0 | 225 | 0 | 0 |

Table 2. Sensitivity analysis to $L_{c}$ for estimated mortality history for Puerto Rico yellowtail snapper. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}_{\boldsymbol{1}}$ | Change Year 1 | $\boldsymbol{Z}_{2}$ | Change Year 2 | $\boldsymbol{Z}_{3}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 260 | 0.55 | 1993.2 | 0.24 | 2005.3 | 0.48 |
| $\mathbf{2 8 0}$ | $\mathbf{0 . 5 6}$ | $\mathbf{1 9 9 2 . 8}$ | $\mathbf{0 . 2 6}$ | $\mathbf{2 0 0 4 . 5}$ | $\mathbf{0 . 4 8}$ |
| 300 | 0.54 | 1992.8 | 0.24 | 2004.5 | 0.41 |
| 320 | 0.48 | 1992.5 | 0.19 | 1996.6 | 0.35 |

Table 3. Sensitivity analysis to $L_{c}$ for estimated mortality history for Puerto Rico hogfish. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}_{\boldsymbol{1}}$ | Change Year 1 | $\boldsymbol{Z}_{\mathbf{2}}$ | Change Year 2 | $\boldsymbol{Z}_{\mathbf{3}}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 260 | 0.33 | 1989.7 | 0.51 | 2001.7 | 0.32 |
| $\mathbf{2 8 0}$ | $\mathbf{0 . 3 6}$ | $\mathbf{1 9 9 2 . 8}$ | $\mathbf{0 . 5 6}$ | $\mathbf{2 0 0 0 . 7}$ | $\mathbf{0 . 3 4}$ |
| 300 | 0.36 | 1993.6 | 0.57 | 2000.3 | 0.36 |
| 320 | 0.37 | 1995.0 | 0.59 | 1999.9 | 0.36 |

Table 4. Sensitivity analysis to $L_{c}$ for estimated mortality history for St. Thomas/St. John queen triggerfish. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}$ |
| :--- | :--- |
| 300 | 1.15 |
| $\mathbf{3 2 0}$ | $\mathbf{1 . 3 4}$ |
| 340 | 1.53 |
| 360 | 1.70 |

Table 5. Sensitivity analysis to $L_{c}$ for estimated mortality history for St. Thomas/St. John spiny lobster. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}_{\boldsymbol{1}}$ | Change Year | $\boldsymbol{Z}_{\mathbf{2}}$ |
| :--- | :--- | :--- | :--- |
| 90 | 0.51 | 1985.7 | 0.76 |
| $\mathbf{1 0 0}$ | $\mathbf{0 . 5 9}$ | $\mathbf{1 9 8 6 . 0}$ | $\mathbf{0 . 9 8}$ |
| 110 | 0.60 | 1986.3 | 0.94 |
| 120 | 0.61 | 1985.7 | 1.10 |

Table 6. Sensitivity analysis to $L_{c}$ for estimated mortality history for St. Croix spiny lobster. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}_{\boldsymbol{1}}$ | Change Year | $\boldsymbol{Z}_{\boldsymbol{2}}$ |
| :--- | :--- | :--- | :--- |
| 80 | 0.35 | 1979.1 | 0.73 |
| $\mathbf{9 0}$ | $\mathbf{0 . 3 5}$ | $\mathbf{1 9 7 9 . 6}$ | $\mathbf{1 . 1 1}$ |
| 100 | 0.74 | 1989.7 | 1.38 |
| 110 | 0.90 | 1990.2 | 1.69 |

Table 7. Sensitivity analysis to $L_{c}$ for estimated mortality history for St. Croix stoplight parrotfish. Bold indicates $L_{c}$ chosen from modal length of length frequency histogram.

| $\boldsymbol{L}_{\boldsymbol{c}}$ | $\boldsymbol{Z}$ |
| :--- | :--- |
| 250 | 1.88 |
| $\mathbf{2 7 0}$ | $\mathbf{2 . 3 1}$ |
| 290 | 2.64 |

Table 8. Derived overfishing limits using the life history point estimates, mean length mortality estimator, per-recruit analyses, and recent catch.

| Stock | Time Period for $\boldsymbol{Z}_{\text {recent }}$ | Estimated $Z_{\text {recent }}$ | M | $\boldsymbol{F}_{\text {recent }}$ | $F_{0.1}$ | $\boldsymbol{F S P R} \%$ \% | Mean catch (pounds) | Time Period of catch | $\begin{gathered} \hline \text { OFL } \\ F_{0.1} \end{gathered}$ | $\begin{array}{r} \text { OFL } \\ \text { F }_{\text {SPR\% }} \text { * } \end{array}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Yellowtail | 2005-2013 | 0.48 | 0.19 | 0.29 | 0.21 | 0.44 | 278,060 | 2005-2014 | 201,354 | 421,884 |
| snapper (PR) <br> Hogfish (PR) | 2001-2013 | 0.34 | 0.16 | 0.18 | 0.11 | 0.17 | 74,931 | 2001-2014 | 45,791 | 70,768 |
| Queen | 1983-2013 | 1.34 | 0.26 | 1.08 | 0.29 | 0.90 | 69,274 | 1998-2014 | 26,406 | 62,249 |
| triggerfish (STT) |  |  |  | $\mathrm{u}=0.66$ | $\mathrm{u}=0.25$ | $\mathrm{u}=0.59$ |  |  |  |  |
| Spiny lobster | 1986-2013 | 0.98 | 0.35 | 0.63 | 0.39 | 0.83 | 88,117 | 1986-2014 | 60,882 | 106,318 |
| (STT) |  |  |  | $\mathrm{u}=0.47$ | $\mathrm{u}=0.32$ | $\mathrm{u}=0.56$ |  |  |  |  |
| Spiny lobster | 1980-2013 | 1.02 | 0.35 | 0.76 | 0.30 | 0.68 | 80,136 | 1980-2014 | 29,279 | 55,736 |
| (STX) |  |  |  | $\mathrm{u}=0.53$ | $\mathrm{u}=0.26$ | $\mathrm{u}=0.49$ |  |  |  |  |
| Stoplight parrot- | 1996-2013 | 2.31 | 0.30 | 2.01 | 0.23 | 0.38 | 90,089 | 1996-2014 | 21,374 | 32,887 |
| fish (STX) |  |  |  | $\mathrm{u}=0.87$ | $\mathrm{u}=0.21$ | $\mathrm{u}=0.32$ |  |  |  |  |

* A spawning potential ratio (SPR) of $40 \%$ was used for St. Thomas/St. John spiny lobster and $30 \%$ for all other stocks.

7. Figures


Figure 1. Length frequency histograms for (a) Puerto Rico yellowtail snapper from handline gear, (b) Puerto Rico hogfish from diving/by hand/spear gear, (c) St. Thomas/St. John queen triggerfish from trap gear, (d) St. Thomas/St. John spiny lobster from trap gear, (e) St. Croix spiny lobster from diving/by hand/spear gear, and (f) St. Croix stoplight parrotfish from diving/by hand/spear gear. Vertical red lines indicate the modal length of the histogram. All length bins are 10 mm .


Figure 2. Annual modal lengths for (a) Puerto Rico yellowtail snapper from handline gear, (b) Puerto Rico hogfish from diving/by hand/spear gear, (c) St. Thomas/St. John queen triggerfish from trap gear, (d) St. Thomas/St. John spiny lobster from trap gear, (e) St. Croix spiny lobster from trap gear, and (f) St. Croix stoplight parrotfish from diving/by hand/spear gear.


Figure 3. Annual length frequency histograms for St. Thomas/St. John spiny lobster from trap gear. Length bins are 10 mm .


Figure 4. Annual length frequency histograms for St. Croix spiny lobster from dive/by hand/spear gear. Length bins are 10 mm .


Figure 5. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for Puerto Rico yellowtail snapper from handline gear, using $L_{c}$ of 280 mm .


Figure 6. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for Puerto Rico hogfish from diving/by hand/spear gear, using $L_{c}$ of 280 mm .


Figure 7. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for St. Thomas/St. John queen triggerfish from trap gear, using $L_{c}$ of 320 mm .


Figure 8. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for St. Thomas/St. John spiny lobster from trap gear, using $L_{c}$ of 100 mm .


Figure 9. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for St. Croix spiny lobster from diving/by hand/spear gear, using $L_{c}$ of 90 mm .


Figure 10. Observed (black) and predicted (red) mean lengths [left] and residuals [right] for St. Croix stoplight parrotfish from diving/by hand/spear gear, using $L_{c}$ of 270 mm .


Figure 11. Distribution of OFL estimates from DLMtool using $F_{0.1}$ (YPR_ML, black) and $F_{S P R \%}$ (SPR30_ML or SPR40_ML, red) for (a) Puerto Rico yellowtail snapper, (b) Puerto Rico hogfish, (c) St. Thomas/St. John queen triggerfish, (d) St. Thomas/St. John spiny lobster (which used $F_{40 \%}$ ), (e) St. Croix spiny lobster, and (f) St. Croix stoplight parrotfish. Vertical lines represent the median of the distribution. Black and red dots represent OFL estimate using mean catch and point estimates for life history for $F_{0.1}$ and $F_{S P R \%}$, respectively.

