1	Genetic population structure of hogfish (Labridae: Lachnolaimus
2	maximus) in the southeastern United States
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19 ABSTRACT

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21 Hogfish (Labridae: Lachnolaimus maximus) stock structure is poorly documented, 22 confounding assessment of this reef fishery species. In this study, genetic structure was examined over a portion of the southeastern U.S. range of the species using a suite of 24 23 24 microsatellite DNA loci. Fin clips from 719 specimens were obtained from multiple geographic areas locations, ranging from Big Bend, Florida to the Carolinas. Genomic 25 proportions of individual hogfish were partitioned into three distinct genetic clusters, 26 geographically delineated as (1) the eastern Gulf of Mexico, (2) the Florida Keys and the 27 28 southeast coast of Florida, and (3) the Carolinas. Clusters 1 and 2 converged along the 29 coastal area west of the Florida Everglades, wherein the intervening samples comprised an 30 admixture of individuals from adjacent clusters. Due to a discontinuity in sample coverage between cluster 2 and 3, the nature and precise location of the genetic break between the 31 Florida Keys and the Carolinas requires further study. The presence of geographically-32 33 limited reproductive exchange in the hogfish leads us to recommend regionally partitioned 34 collections and analyses of life-history and fishery data for stock assessments. 35

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37 INTRODUCTION

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39 A fundamental challenge to managing sustainable fisheries involves aligning biological 40 evidence of stock structure with fishing and management sectors for the purpose of monitoring, assessment, and regulatory actions (Cadrin et al. 2014). This is particularly 41 42 challenging around the Florida peninsula, where several zoogeographic regions overlap state (Florida) and federal boundaries, and fall under the jurisdiction of two federal fishery 43 44 management councils (South Atlantic and Gulf of Mexico). In terms of zoogeography, there 45 is a strong environmental gradient along (north-south) and between (east-west) coasts, resulting in distinctive faunal breaks at Cape Romano (west coast, Gulf of Mexico) and Cape 46 47 Canaveral (east coast, Atlantic Ocean; Briggs and Bowen 2012). Subspecies are frequently 48 recognized between the east and west coasts of Florida (Bowen and Avise 1990) and several coastal and marine species are considered to have distinct stocks on the respective coasts 49 50 (Tringali and Bert 1996, Gold et al. 2002, McBride in press).

51

52 Many fishery species in the region remain data-poor with regard to stock structure. One such example is the hogfish (Labridae: Lachnolaimus maximus), a long-lived reef species that 53 54 occurs in temperate to tropical waters of the western Atlantic Ocean and Gulf of Mexico (Claro et al. 1989, McBride and Richardson 2007, McBride et al. 2008). The species 55 56 supports a modest commercial fishery in the southeastern United States, and is a highly 57 valued target for recreational divers and spear fishers (McBride and Murphy 2003, FWC 2013, NMFS 2013). To date, no information on hogfish stock structure exists other than 58 general behavioral and life history patterns and a preliminary genetic survey in the eastern 59 60 Gulf of Mexico. Like most marine fishes, female hogfish broadcast eggs for external fertilization, which facilitates dispersal of the propagules away from the spawning site (Colin 61

1982). The planktonic larval duration lasts 3-5 weeks (Colin 1982, Victor 1986), which is fairly modest among reef fishes (Victor 1986, Soeparno et al. 2012), and does not imply extensive mixing of genotypes between ocean basins. Moreover, hogfish spawn in stable, site-specific harems throughout their range (i.e., they do not migrate or periodically aggregate to spawn) and there is specific evidence of reciprocal onshore larval dispersal and offshore movements by subadults (Muñoz et al. 2010, Collins and McBride 2011). These behaviors create demographic structure that may interact with stock structure.

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70 Life history differences for hogfish (e.g., maximum age, maximum size and size and 71 fecundity) exist between the eastern Gulf of Mexico and south Florida (McBride et al. 2008) 72 and even within south Florida (McBride and Richardson 2007). These life history patterns 73 were interpreted as a result of high fishing rates in more accessible habitat, and although they affect vital rates within each region, they have not been linked to underlying biological stock 74 75 structure (McBride and Richardson 2007, McBride et al. 2008, McBride 2014). Questions 76 remain, most recently as a request to review hogfish stock structure and unit stock definitions as part of the 2014 'Southeastern U.S. Hogfish Benchmark Assessment' 77 78 (http://www.sefsc.noaa.gov/sedar/, SEDAR 37).

79

This paper introduces new genetic data to determine if hogfish stock structure exists within the coastal waters of the southeastern U.S., and if so, where the stock boundaries occur. Because genetic differentiation within marine populations can be difficult to detect, this investigation utilized microsatellite DNA markers, which possess a number of attributes that are well suited for such work relative to other marker types (e.g., allozymes and mitochondrial DNA) (Mariani and Bekkevold 2014, Antoniou and Magoulas 2014). Microsatellites are nuclear encoded, co-dominant markers that have characteristically high 87 mutation rates and, thus, a high degree of allelic variation. These loci are scattered 88 throughout the genome and can be influenced independently by recombination, selection and 89 drift; thus, each locus is expected to have its own slightly different genealogical history. 90 Combining many loci makes a genomic sampling increasingly representative of the above genetic processes, and provides a more robust method for investigating gene flow and 91 92 population connectivity (Hedrick 1999; Kalinowski 2002, 2005; Wilson and Rannala 2003). Here we apply microsatellite loci previously isolated for hogfish and optimized for routine 93 94 assay (Seyoum et al. 2012) to survey specimens collected from Big Bend, Florida to the 95 Carolinas.

96

97 MATERIALS AND METHODS

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99 Sample collection and DNA extraction

100 Study specimens (n = 719) were collected through intercepting recreational and commercial 101 spear fishers or during directed research trips by biologists from the Florida Fish and Wildlife 102 Conservation Commission's Fish and Wildlife Research Institute (FWC FWRI). Fin clips were removed and preserved in 70% ethanol. Total DNA was isolated from approximately 103 104 500 mg of fin clip tissue using PUREGENE DNA isolation kits (Gentra Systems, INC., 105 Minneapolis Minnesota) and rehydrated in 50 µl of deionized water. 106 107 Collection locations were subdivided into nine geographic areas, referred to hereafter as 108 'samples' (Figure 1). These study samples were established predominantly by latitude and

109 coast (west [Gulf of Mexico], east [Atlantic Ocean]) to delineate areas corresponding to

110 recognized faunal breaks, major estuaries, and, on a broader scale, management jurisdictions.

111 For example, faunal breaks occur at Cape Romano between samples 5 and 6, Cape Sable

112 between samples 6 and 7, and Cape Canaveral between samples 8 and 9. Considerable estuarine flow onto the continental shelf occurs from the Suwannee River between samples 1 113 114 and 2, Tampa Bay (sample 3) and Charlotte Harbor (sample 4). In terms of management 115 jurisdictions, the Gulf of Mexico Fishery Management Council regulates federal fisheries throughout the Gulf of Mexico (samples 1-6); the South Atlantic Fishery Management 116 117 Council regulates fisheries in federal waters from the Carolinas to the Florida Keys (samples 7-9); and Florida regulates state fisheries inside 9 miles in the Gulf of Mexico and inside of 3 118 119 miles in the Atlantic (samples 1-8).

120

121 Microsatellite genotyping

122 Specimens were genotyped using 24 of 29 microsatellite markers identified in Seyoum et al.

123 (2012); (Lmax11, Lmax14, Lmax15, Lmax24, and Lmax31 markers were not used).

124 Multiplex PCR amplifications were carried out in an Eppendorf thermal cycler containing 50-

125 100 ng of total DNA, 10 µl 50uM of dNTP mix, 0.25 µl of 0.1mg/ml BSA, a combination of

three optimally selected primers of three loci with each forward primer labeled with different

127 fluorescent dye, 5 μl of *Taq* Polymerase Buffer (10×) containing 15mM MgCl₂ (Promega)

and 1.25 units of *Taq* polymerase (Promega). The reaction profile was 94°C for 2 min, $35 \times$

129 (94°C for 35 s, 55°C for 35 s, 72°C for 35 s) and final extension at 72°C for 30 min.

130 Fragments were visualized on an ABI 3130 XL genetic analyzer and genotyped using

131 GeneMapper software (version 4.0; Applied Biosystems, Inc.). For fragment assays, we used

132 Gene Scan-500 ROX-labeled size standard (Applied Biosystems, Inc. Foster City, CA).

133

134 Data analysis

135 GENEPOP data files were generated from fragment sizes recorded using the Microsatellite

136 Marker Toolkit add-on (version 3.1.1; Park 2001): <u>http://animalgenomics.ucd.ie/sdepark/ms-</u>

137 toolkit/); these were converted to other formats as needed using the conversion tool PGDspider (version 2.0.1.9; Lischer and Excoffier 2012). Pairwise genetic distances (F_{ST}) 138 139 between samples (Weir and Cockerham 1984) were estimated with 10,000 permutations in 140 the program GENETIX (Belkhir et al. 2000). Departures from Hardy-Weinberg equilibrium (HWE) were determined by using GENEPOP (version 3.4; Raymond and Rousset 1995, 141 142 Rousset 2008). Sequential Bonferroni corrections were applied to multiple tests of hypotheses (Rice 1989). Observed (H_0) and expected heterozygosity $(H_e, with and without a$ 143 144 bias correction), averaged over all loci, were obtained from GENETIX (Belkhir et al. 2000). 145 Null allelism was investigated with the randomization test of Guo and Thompson (1992) and the U-test statistic of Raymond and Rousset (1995), using the program ML-NullFreq 146 147 (http://www.montana.edu/kalinowski/software.htm). For each locus, microsatellite variation 148 was quantified in terms of genetic diversity, number of alleles, and allelic richness (a diversity measure that corrects for differences in sample size) using the program FSTAT 149 150 (version 2.9.3.2; Goudet 2001). Chi-square tests were performed to determine if samples 151 differed significantly with respect to the above standard genetic measures. 152

153 Phenetic clustering

154 Corrected and uncorrected F_{ST} genetic distances (Latter 1972) were generated from allele

155 frequencies and Unweighted Pair Group Method with Arithmetic Mean (UPGMA)

156 phenograms based on these distances were reconstructed using the Software POPTREE2

157 (Takezaki, Nei, and Tamura 2010) and visualized using the software FIGtree

158 (<u>http://tree.bio.ed.ac.uk/software/figtree/</u>). Statistical significance of relationships was

assessed via 1000 bootstrap replications. Adopting patterns observed in the phenogram,

160 samples were categorized into geographic clusters for further analysis of genetic structure.

162 Genetic structure

163 Genetic structure among the study samples was examined using three analytical approaches. The first was based on the analysis of molecular variance (AMOVA) as implemented in the 164 165 program ARLEQUIN (version 3.5.1.3; 100,000 permutations, Excoffier and Lischer 2010). AMOVA follows an *a priori* hierarchical approach in which correlations among genotypes at 166 167 various levels are partitioned as F-statistics. The proportion of variation among groups (F_{CT}), within groups (F_{SC}), and within samples (F_{ST}) was computed and the φ statistic assessed by 168 the permutation method of Excoffier et al. (1992). Initially, the *a priori* hierarchical structure 169 170 analyzed was based on the results of above phenetic reconstruction. For comparison, sample 171 groupings other than that indicated by the phenogram were also considered. 172 173 The second analytical approach was based on the Bayesian population-assignment algorithm as implemented in the program STRUCTURE (version 2.3.4; Pritchard et al. 2000). With 174 175 this algorithm, individuals were probabilistically and proportionally assigned to one or more 176 genetic clusters (K) in a manner that minimized Hardy-Weinberg and linkage disequilibrium 177 among their multilocus genotypes. For K = 1 through 9, ten simulations were conducted using 2E6 Markov-Chain Monte Carlo (MCMC) replicates following a 1E6 burn-in period. 178 179 We adopted the admixture model and the independent allele frequency option to minimize 180 the chance of overestimating the number of clusters present in the data (Pritchard et al. 2009). 181 We used STRUCTURE HARVESTER (version 0.6.93; Earl and vonHoldt 2012) with each 182 of the above replicate runs to compute the *ad hoc* statistics Ln(K) and ΔK in order to determine the most plausible 'base' value of K clusters (i.e., the upper-level hierarchy). Ln 183 (K) denotes the log probability of the data at a given modeled K value; ΔK is based on the 184

185 rate of change in Ln(K) between successively modeled K values. Simulation studies

186 (Evanno et al. 2005) have shown that ΔK provides the most accurate indication of genetic

structure under a variety of modeling conditions. We then used CLUMPP (version 1.1.2;

188 Jakobsson and Rosenberg 2007) to determine the optimal alignment for replicate analyses

and mean genomic membership coefficients across replicate runs for samples and

190 individuals. The sample coefficients were plotted using EXCEL (Microsoft) and the 719

191 individuals using DISTRUCT (version 1.1; Rosenberg 2004) and visualized using the

192 program Ghost View (http://pages.cs.wisc.edu/_ghost).

193

Finally, to determine if genetic relationships among the samples conformed to a pattern of
isolation-by–distance gene flow (Malécot 1955; Wright 1943), we computed the mantel
correlation coefficient R between genetic distance (F_{ST}) and geographic distance (Km) using
the program GenAlEx (version 6.5; Peakall and Smouse 2006, 2012). Significance of the
correlation coefficient was tested via 9000 random permutations.

200 **RESULTS**

201

202 Standard genetic measures and distances

203 Significant heterozygote deficiencies were sporadically detected at three loci (Lmax4,

Lmax29 and Lmax 35) in one to five samples when biased estimates of expected

205 heterozygosities were considered; however, no deviations were observed when unbiased

206 estimates were considered. Presumptive frequencies of null alleles at those loci exhibiting

207 heterozygote deficits ranged from 0.17 to 0.24. There were no significant differences among

samples in the mean values of standard genetic measures (Table 1). For all loci, a total of

209 350 unique alleles were identified and of these, an average of 216 alleles occurred in each

sample. Over the 36 possible pairwise comparisons, 30 sample pairs had genetic distances

211 that were significantly greater than zero (P < 0.05).

212

214

213 Phenetic clustering

oriented but poorly supported clusters. Cluster 1 included specimens collected in the eastern Gulf of Mexico from the Florida Panhandle (Big Bend) to Naples (samples 1 - 5). Cluster 2 included specimens collected from the Florida Keys and along the southeastern coast of Florida (samples 7 - 8). Cluster 3 included specimens collected from the Carolinas (sample 9). Significant bootstrap support was observed only for the branch separating sample 9 (Cluster 3) from the remaining samples (clusters). The position of the sample 6 (Everglades region) was unresolved in the reconstruction, fluctuating between Cluster 2 (Figure 2A;

UPGMA reconstructions based on F_{ST} distances (Figure 2) revealed three geographically

222 uncorrected F_{ST}) and Cluster 1 (Figure 2B; corrected F_{ST}).

223

224 Analysis of Molecular Variance

In the AMOVA, 98.17% of the variation occurred within samples and 1.83% among samples. The overall F_{ST} value of 1.83 (*P*<0.0001) was indicative of significant differentiation among samples due to the presence of spatial structure at both regional and local scales. The highest among-group variance in the AMOVA (Table 2) was observed when samples were grouped according to the phenetic clusters shown in Figure 2B. However, translocation of sample 6 from Cluster 1 to Cluster 2 only slightly reduced the among-group variance.

232

233 Pairwise F_{ST} values between uncorrected (and corrected) F_{ST} phenetic clusters, respectively,

234 were 0.015 (0.016) between Clusters 1 and 2, 0.027 (0.023) between 1 and 3, and, 0.030

(0.039) between 2 and 3 (Table 2). All pairwise F_{ST} comparisons among the clusters were

highly significant (P<0.001). Overall, the F_{ST} comparisons and the AMOVA confirmed the

arranged with the least genetic differentiation between Clusters 1 and 2, and the
differentiation between Clusters 2 and 3 exceeding that observed between the geographically
disjunctive Clusters 1 and 3. Thus interestingly, the observed pattern of genetic

presence of three geographically based clusters, but these appeared to be hierarchically

241 differentiation did not conform to isolation-by-distance expectations over the studied range.

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237

243 **Bayesian population assignment**

In the absence of a methodology to determine whether values of Ln(K) statistically differ,

inferences were derived herein by evaluating replicate likelihoods and resultant ΔK statistics

for different values of K. Ln(K) values increased quickly from K = 1 to K = 2, and somewhat

less quickly from K = 2 to K = 3, before reaching a plateau at successive values (Figure 3A).

248 The largest value of ΔK occurred for K = 2 ($\Delta K = 281$) and, secondarily, for K = 3 ($\Delta K =$

118) (Figure 3B). At the base hierarchical level of K=2, members of the phenetically

250 delineated Clusters 1 and Cluster 3 were predominantly conjoined within a single Bayesian

251 cluster whereas members of phenetic Cluster 2 were predominantly assigned to a discrete

252 Bayesian cluster (Figure 4A). At the secondary hierarchical level of K = 3, members of

253 phenetic Clusters 1 and 3 were assigned to separate Bayesian clusters (Figure 4B). The

254 CLUMMP analysis indicated that samples 5 and 6, respectively, were admixtures of

Bayesian Clusters 1 and 2, exhibiting graduated mean genomic proportions (Figure 4C).

256

257 Mantel Test

There was only a sufficient number of paired comparisons within Cluster 1 (Big Bend to Everglades) to perform a within-cluster Mantel test. There was no significant correlation between genetic and geographic distances (p = 0.125, R = 0.334) (Figure 5).

262 In summary, the phenetic clustering, AMOVA, and Bayesian population-assignment analyses elucidated a concordant pattern of genetic structure within the studied range. Three 263 264 geographically based clusters were delineated as (1) the eastern Gulf of Mexico, (2) the 265 Florida Keys and the southeastern coast of Florida, and (3) the Carolinas (Figure 6). There was no indication of isolation-by-distance gene flow within the eastern Gulf or over the 266 267 sampled range. The area west of the Florida Everglades (sample 6) appeared to serve as a genetic break between the respective stocks. Due to a discontinuity in sample coverage, the 268 269 nature and precise location of the apparent genetic break between the Florida Keys and the 270 Carolinas requires further study.

271

272 **DISCUSSION**

273

At the outset of this study, three genetic stocks of hogfish could not be confidently predicted, 274 275 but neither were they unexpected at this landscape scale. Capes Romano and Canaveral have 276 previously been identified as faunal breaks, associated with major shifts in marine animal 277 communities. Briggs and Bowen (2012) identify Cape Romano as separating the Gulf of Mexico from the Caribbean province, and Cape Canaveral as separating the Carolinian 278 279 province from the Caribbean province. Strong genetic differences between the two Florida 280 coasts are known for several inshore and nearshore fishes including common snook, 281 (Centropomus undecimalis) (Tringali and Bert 1996), red drum (Sciaenops ocellatus) (Gold 282 et al. 1999; Seyoum et al. 2000; Gold and Turner 2002), and spotted seatrout (Seyoum et al. 2014). In contrast, only weak genetic differences, if any, have been found between Florida 283 coasts for the reef fishes such as red grouper (Epinephelus morio) and scamp (Mycteroperca 284 285 phenax) (Zatcoff et al. 2004), vermilion snapper (Tringali and Higham 2007), and for the pelagic king mackerel (Scomberomorus cavalla) (Gold et al. 2002). Based upon existing 286

data, hogfish stock structure and levels of population connectivity in the southeastern United
States are generally more similar to inshore and nearshore fishes than to offshore reef fishes
or large coastal pelagics.

290

Planktonic larval duration is a good first approximation of the potential for gene flow in 291 292 marine systems (Waples 1987, Purcell et al. 2006, Soeparno et al. 2012). Hogfish eggs and larvae are transported away from the reef habitats where spawning occurs, and juveniles 293 294 settle within shallow, nearshore grass beds (Davis, 1976; Colin 1982). They have a modest 295 planktonic larval duration lasting approximately 35 days before strong benthic orientation 296 begins (Colin 1982, Victor 1986). Post hatch larvae are found under floating material, and 297 may also form bubbles around themselves while floating near the surface (Colin 1982), 298 suggesting that distribution of larvae would be influenced more heavily by surface currents 299 than mid-water or benthic transport.

300

301 Nonetheless, the hydrodynamics around Florida are complex enough that predicting the 302 outcome of the average flow, in the average direction, for the average planktonic period is too 303 simplistic. Along the west Florida shelf, current flows are weakest near Cape Romano, which 304 may present a barrier to gene flow between clusters 1 and 2; however, current direction above 305 the shelf also shifts from southerly to northerly throughout the year, and at least near the shelf 306 break, a strong looping current interacts periodically with local wind forces above the shelf 307 (He and Weisberg 2003). The peak spawning period for hogfish in the eastern Gulf of 308 Mexico (March - May; McBride et al. 2008; Collins and McBride, in review) occurs at a time when the currents shift from a predominant southeast flow to a northwestern flow (Liu 309 310 and Weisberg 2005, 2012). However, spawning occurs to some extent in almost all months of the year (except September; Collins and McBride, 2011, in review). Spawning throughout the 311

312 year would subject larvae to a diverse (and difficult to predict) set of conditions, both 313 physical (e.g., seasonal shifts in currents, temperature) and biological (e.g., prev availability 314 through seasonal plankton blooms) that affect dispersal vectors and survival rates. The Loop 315 Current in the Gulf of Mexico is an upstream portion of the Gulf Stream and impacts portions of the west Florida shelf to differing degrees throughout the year. The Loop Current 316 317 presumably occurs beyond the range of hogfish spawning activity on the west Florida shelf (Cluster 1; spawning < 70 m; Collins and McBride, in review), but the effects of this major 318 319 current and its associated eddies are difficult to assess since many regional ichthyoplankton 320 surveys report taxa to only the family level (e.g., Labridae; Richards et al. 1993). Along the 321 Atlantic side of the Florida Keys, the dominant current flow is to the east, fed by the Florida 322 Current that flows through the Florida Strait and into the Gulf Stream. This current flow 323 suggests that the majority of larvae produced south of the Everglades region would be transported into the Keys or Atlantic Ocean with very little larval transport north along the 324 325 west Florida shelf. Models of regional hydrodynamics combined with the increased 326 information available regarding hogfish distribution and spawning could simulate larval 327 dispersal and settlement around Florida. Elsewhere, biological data are lacking; for example, courtship of hogfish has been observed off of North Carolina (Parker 2000), but the timing 328 329 and extent of spawning in this region (Cluster 3) has not been described.

330

Adult behavior may also play an important role in determining connectivity and population structure (Frisk et al., in press) All hogfish mature as a female first, but if they live long enough they will eventually change sex (McBride and Johnson 2007). A single male (which may have been previously a female in the same harem) will spawn daily with multiple females in harems containing as many as 15 individuals (Colin 1982, McBride and Johnson 2007, Muñoz et al. 2010; Collins and McBride 2011). Mature hogfish have relatively strong site fidelity to the same reef habitats for months and even years (Colin 1982, Lindholm 2006,
Muñoz 2010). However, fish less than 2 years old have rarely been observed from offshore
habitats (Collins and McBride 2011) and because juveniles settle in shallow, inshore habitats,
it is assumed that they gradually move offshore with growth. Although natal homing is not
predicted for hogfish, it has been demonstrated for other reef fish (Paris et al. 2013); at the
very least, such reciprocal ontogenetic movements demonstrated by hogfish focus gene flow
in an inshore-offshore direction rather than along-shore direction.

344

345 Demographic differences in hogfish between or within regions of Florida have been noted 346 (McBride and Richardson 2007, McBride et al. 2008, Collins and McBride 2011). In the 347 central eastern Gulf of Mexico, adult size, longevity, and fecundity differ between deep and 348 shallow water fish across the shelf (Collins and McBride 2011), unrelated to genetic structure (Seyoum et al. 2012). Although demographic differences also exist between the 349 350 Gulf of Mexico and south Florida (McBride and Richardson 2007, McBride et al. 2008), 351 which can now be said to be genetically different, it is premature to assume these life history 352 differences are a result of genetic differences. The effect of fishing on size and at age, maturity, fecundity, and harem stability is very likely playing a large role in the 353 354 demographics of hogfish throughout their range (McBride et al. 2008, Muñoz et al. 2010). In 355 addition to fishing pressure, the potential effects of red tide on recruitment rates should be 356 considered together with this new information on genetic stock structure around Florida 357 (McBride and Richardson 2007, Collins and McBride 2011).

358

359 If the inferred close relationship between Clusters 1 and 3 is real, the genetic forces 360 responsible are not readily apparent. It could be possible that hogfish in Clusters 1 and 3 are 361 connected to some degree by gene flow to a larger Caribbean stock via the Gulf Stream 362 current (CITE), with those in Cluster 2 being spatially isolated due to localized hydrodynamics and/or spawning behaviors. Alternatively, it is possible that for unknown 363 demographic reasons genetic drift has had a greater relative effect on the neutral genetic 364 365 composition of Cluster 2. Finally, it is possible that temperature, or some other correlate along a latitudinal cline, could be driving selection for particular genotypes. Temperatures 366 367 are high (> 25 C) and relatively homogeneous around Florida during the summer (June-September), but the temperatures over the western Florida shelf are much colder than in south 368 369 Florida (Cluster 2) during the rest of the year (He and Weisberg 2003). Such a latitudinal 370 cline could result in the closer genetic relationship between Clusters 1 and 3, where yet 371 unknown environmental effects may be selecting for a more similar genotype. Additional 372 study is needed to address these hypotheses.

373

Application of our findings to monitoring, assessment, and regulatory actions should not be 374 375 problematic. Landings and fishing effort in Florida are already collected at the county level 376 (FWC 2013), and the Florida Keys are contained with a single county (Monroe). Federal 377 fisheries data can also be separated by coast (e.g., NMFS 2013) and we predict that regional 378 assessments will further demonstrate differences in fishing pressures between regions: fishing 379 pressure within Cluster 1 is high (annual landings data reference for the Gulf, NMFS 2013) 380 but the pressure is more spread out due to the nature of the habitat and the expansive west 381 Florida shelf. Fishing pressure is also high in Cluster 2, where it is condensed into a much 382 smaller area that is more accessible to divers. In both regions, cluster 1 and 2, hogfish residing in shallow depths are cropped out of the population as soon as they reach legal sizes, 383 particularly in the most (human) populated areas (McBride and Richardson 2007). Cluster 3 384 385 presumably has lower fishing pressure; there are commercial and recreational fisheries in this area, but the travel time, distance from shore, weather patterns and lower human population 386

387 density likely reduce fishing pressure for this species off of the Carolinas. Regional assessments that can quantify fishing effort and landings have been challenging in the past 388 389 (Kingsley 2004). As an alternative, Ault et al. (2005) promoted a size-based approach to 390 addressing the data poor nature of the hogfish assessment, and Collins and McBride (2011) underscore considerations of spatial demographic structure. The genetic data presented 391 392 herein strongly suggest that regional assessments are warranted for the eastern Gulf of 393 Mexico, south Florida, and a broad but as yet undefined area from northeast Florida to the 394 Carolinas.

395

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2	Figure 1.	Hogfish (Lachnolaimus maximus) specimens were grouped by collection location into
3		samples, where 1 = Big Bend, 2 = Nature Coast and Middle Grounds, 3 = Tampa Bay, 4
4		= Sarasota, 5 = Naples, 6 = Everglades, 7 = Florida Keys, 8 = East Florida, 9 = Carolinas.
5		
6	Figure 2.	Phenograms based on (A) uncorrected and (B) corrected genetic distances (F_{ST}) between
7		each pair of the 9 hogfish samples as identified through the Unweighted Pair Group
8		Method with Arithmetic Mean (UPGMA: 1000 bootstrap replications) hierarchical
9		clustering method using the program POPTREE2.
10		
11	Figure 3.	(A) Mean likelihood L (posterior probability) and (B) ΔK as determined from 10
12		replicates of each value of K (from 1 to 9). Statistics are from STRUCTURE analysis of
13		the 9 hogfish collection samples.
14		
15	Figure 4.	Genetic population structure among hogfish samples according to posterior probability
16		assignment produced by the STRUCTURE analysis of 24 polymorphic microsatellite loci
17		$(4.5 \times 10^6 \text{ burn-in, and } 9.0 \times 10^6 \text{ replications})$: The CLUMMP output from 10 replicates for
18		719 individuals and the 9 samples; (A) for the highest modal value (number of genetic
19		clusters) at $K=2$, (B), for the next modal value of $K=3$ shown in graphic display (hogfish
20		population structure) using the program DISTRUCT, and (C) proportional values of each
21		cluster for each of the 9 samples. Figures are demarcated by vertical lines.
22		

24		
25	Figure 5.	Mantel tests examining correlations between pairwise genetic distance (F_{ST}) and
26		geographic distance (km) within Cluster 1 from Big Bend (sample 1) to the Everglades
27		(sample 6).
28		
29	Figure 6.	Geographic boundaries of the three hogfish distinct clusters based on the genetic
30		population structure, as detected using 24 microsatellite loci. Dotted lines indicate
31		flexible demarcations of clusters potentially expanding in the direction indicated by
32		arrows. The region of gene flow restriction (genetic break) between clusters 1 and 2 is
33		identified by a circular delineation.
34		

Table 1. Standard genetic measurements averaged over all loci for each of the nine samples of hogfish. Collection areas for samples appear in Figure 1. In the last column, numbers in parenthesis represent sample averages.

	Samples									
	1	2	3	4	5	6	7	8	9	Total
No. of specimens	119	71	88	24	22	70	191	32	102	719
Genetic diversity	0.63	0.65	0.65	0.63	0.64	0.67	0.67	0.68	0.61	0.65 (0.65)
Total no. of alleles sampled	246	223	237	165	161	238	296	203	174	350 (216)
Average no. of alleles per locus	10.3	9.3	9.9	6.9	6.7	9.9	12.3	8.5	7.3	14.6 (9.0)
Allelic richness	9.5	9.0	9.7	9.6	6.3	9.7	12.2	8.2	7.1	14.4 (9.0)
Observed heterozygosity	0.57	0.59	0.60	0.61	0.59	0.63	0.63	0.63	0.58	0.60 (0.60)
Unbiased observed heterozygosity	0.63	0.65	0.65	0.63	0.64	0.66	0.67	0.68	0.61	0.66 (0.65)
Expected heterozygosity	0.63	0.64	0.64	0.62	0.63	0.66	0.66	0.67	0.60	0.66 (0.64)

1Table 2. Analysis of molecular variance (AMOVA) based on 24 microsatellite loci genotypes of2719 hogfish specimens collected from Big Bend, Florida, to the Carolinas. Numbers joined by3hyphens are samples combined into clusters based on the corrected and uncorrected F_{ST} 4distances. Number codes as in Figure 1. Only the values of the F_{CT} among the three-cluster and5the F_{ST} between the two-cluster analyses are given below.

Among	three-cluster co	mponent	Obse	rved partitio	<u>n</u>	
1 st Cluster	2 nd Cluster	3 rd Cluster	Variance	% Total	F _{CT}	Phenetic cluster from
1-2-3-4-5-6	7-8	9	0.14768	2.07	0.0208	Corrected F _{ST}
1-2-3-4-5	6-7-8	9	0.14331	2.02	0.0202	Uncorrected F _{ST}
Among	two-cluster cor	nponent	<u>Obs</u>	erved partiti	on	
1 st Cluster	2 nd Cluster	3 rd Cluster	Variance	% Total	F _{ST}	Phenetic cluster from
1-2-3-4-5-6	7-8		0.11203	1.56	0.0156	Corrected F _{ST}
1-2-3-4-5-6		9	0.16073	2.30	0.0230	Corrected F _{ST}
	7-8	9	0.28595	3.93	0.0393	Corrected F _{ST}
1-2-3-4-5	6-7-8		0.10629	1.48	0.0149	Uncorrected F _{ST}
1-2-3-4-5		9	0.17683	2.56	0.0256	Uncorrected F _{ST}
	6-7-8	9	0.22038	3.01	0.0301	Uncorrected F _{ST}

6

Probability of finding a more extreme variance component by chance alone (1000 permutations)
<0.001 in all the F_{CT} and F_{ST} statistic.

9





(A) Uncorrected F_{st}

(B) Corrected F_{st}













