# SEAMAP Reef Fish Video Survey: Relative Indices of Abundance of Greater Amberjack 

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SEAMAP Reef Fish Video Survey: Relative Indices of Abundance of Greater Amberjack May 2013<br>Matthew D. Campbell, Kevin R. Rademacher, Paul Felts, Brandi Noble, Michael Felts, and Joseph Salisbury Southeast Fisheries Science Center<br>Mississippi Laboratories,Pascagoula, MS

## Introduction

The primary objective of the annual Southeast Area Monitoring and Assessment Program (SEAMAP) reef fish video survey is to provide an index of the relative abundances of fish species associated with topographic features (e.g reefs, banks, and ledges) located on the continental shelf of the Gulf of Mexico (GOM) from Brownsville, TX to the Dry Tortugas, FL (Figures 1 and 2). Secondary objectives include quantification of habitat types sampled (video and side-scan), and collection of environmental data throughout the survey. Because the survey is conducted on topographic features the species assemblages targeted are typically classified as reef fish (e.g. red snapper, Lutjanus campechanus), but occasionally fish more commonly associated with pelagic environments are observed (e.g. hammerhead shark, Sphyrna lewini). The survey has been executed from 1992-1997, 2001-2002, and 2004-2012 and historically takes place from May August. The 2001 survey was abbreviated due to ship scheduling, during which, the only sites that were completed were located in the western Gulf of Mexico. Types of data collected on the survey include diversity, abundance (minimum count), fish length, habitat type, habitat coverage, and bottom topography. The size of fish sampled with the video gear is species specific however greater amberjack sampled over the history of the survey had fork lengths ranging from 101.0 2065.0 mm , and mean annual fork lengths ranging from $571.8-759.9 \mathrm{~mm}$. Age and reproductive data cannot be collected with the camera gear but beginning with the 2012 survey, a vertical line component will be coupled with the video drops to collect hard parts, fin clips, and gonads.

## Methods

## Sampling design

Total reef area available to select survey sites from is approximately $1771 \mathrm{~km}^{2}$, of which $1244 \mathrm{~km}^{2}$ is located in the eastern GOM and $527 \mathrm{~km}^{2}$ in the western GOM. The large size of the survey area necessitates a two-stage sampling design to minimize travel times between stations. The first-stage uses stratified random sampling to select blocks that are 10 minutes of latitude by 10 minutes of longitude in dimension (Figures 1 and 2). The block strata were defined by geographic region (4 regions: South Florida, Northeast Gulf, Louisiana-Texas Shelf, and South Texas), and by total reef habitat area contained in the block (blocks $\leq 20 \mathrm{~km}^{2}$ reef, block $>20 \mathrm{~km}^{2}$ reef). There are a total of 7 strata. A 0.1 by 0.1 mile grid is then overlaid onto the reef area contained within a given block and the ultimate sampling sites (second stage units) are randomly selected from that grid.


Figure 1. SEAMAP reef fish video survey sample blocks located in the eastern Gulf of Mexico.


Figure 2. SEAMAP reef fish video survey sample blocks located in the western Gulf of Mexico.

## Data reduction

Various limitations either in design, implementation, or performance of gear causes limitations in calculating minimum counts and are therefore dropped from the design-based indices development and analysis as follows. In 1992, each fish was counted every time it came into view over the entire record time and the total of all these counts was the maximum count. Maximum count methodologies are not preferred and the 1992 video tapes were destroyed during Hurricane Katrina and cannot be re-viewed, so 1992 data is excluded from analyses (unknown number of stations). The 2001 survey was abbreviated due to ship scheduling, during which, the only sites that were completed were located in the western GOM. Because of the spatial imbalance associated with data gathered in 2001, that entire year has been dropped ( 80 total sites). Stratum 1 (South Florida) and stratum 7 (S. Texas) are blocks that contain very little reef and were not consistently chosen for sampling and were also dropped ( 184 total sites). Occasionally tapes are unable to be read (i.e. organisms cannot be identified to species) for the following reasons including: 1) camera views are more than $50 \%$ obstructed, 2 ) sub-optimal lighting conditions, 3) increased backlighting, 4) increased turbidity, 5) cameras out of focus, 6) cameras failed to film. In all of these cases the station is flagged as 'XX' in the data set and dropped (190 total sites). Sites that did not receive a stratum assignment are also dropped (62). By these criteria the data set is reduced 4744 down to 4228 sites analyzed.

## Gear and deployment

The SEAMAP reef fish survey has employed several camcorders in underwater housings since 1992. Sony VX2000 DCR digital camcorders mounted in Gates PD150M underwater housings were used from 2002 to 2005 and Sony PD170 camcorders during the years 2006 and 2007. In 2008 a stereo video camera system was developed and assembled at the NMFS Mississippi Laboratories Stennis Space Center Facility and has been used in all subsequent surveys. The stereo video unit consists of a digital stereo still camera head, digital video camera, CPU, and hard drive mounted in an aluminum housing. All of the camcorder housings we have used were rated to a maximum depth of 150 meters while the stereo camera housings are rated to 600 meters. Stereo cameras are mounted orthogonally at a height of 50 cm above the bottom of the pod and the array is baited with squid during deployment.

At each sampling site the stereo video unit is deployed for 40 minutes total, however the cameras and CPU delay filming for 5 minutes to allow for descent to the bottom, and settling of suspended sediment following impact. Once turned on, the cameras film for approximately 30 minutes before shutting off and retrieval of the array. During camera deployment the vessel drifts away from the site and a CTD cast executed, collecting water depth, temperature, conductivity, and transmissivity from the surface to the maximum depth. Seabird units are the standard onboard NOAA vessels however the model employed was vessel/cruise dependent.

## Video tape viewing

One video tape from each station is selected for viewing out of four possible. If all four video cameras face reef fish habitat and are in focus, tape selection is random. Videos are viewed for twenty minutes starting from the time when the view clears from suspended sediment. Viewers identify, and enumerate all species to the lowest taxonomic level during the 20 minute viewable segment. From 1993-2008 the time when each fish entered and left the field of view was recorded a procedure referred to as time in - time out (TITO) and from these data a minimum count was calculated. The minimum count is the maximum number of individuals of a selected
taxon in the field of view at one instance. Each 20 minute video is evaluated to determine the highest minimum count observed during a 20 minute recording. The 2008-2011 digital video allows the viewer to record a frame number or time stamp of the image when the maximum number of individuals of a species occurred, along with the number of taxon identified in the image but does not use the TITO method. Both the TITO and current viewing procedure result in the minimum count estimator of relative abundance. Minimum count methodology is preferred because it prevents counting the same fish more than once and represents the conservative maximum number of fish that were at a location at one point in time.

## Fish length measurement

Beginning in 1995 fish lengths were measured from video using lasers attached on the camera system with known geometry. However, the frequency of hitting targets with the laser is low and precluded estimating size frequency distributions. Additionally, the same fish can be measured more than once at a given station. So, the lengths measured provide the range of sizes observed. The stereo cameras used in 2008-2010 allow size estimation from fish images. The Vision Measurement System (Geometrics Inc.) was used to estimate size of greater amberjack. We estimated a length frequency distribution by weighting station length frequencies by station Minimum Counts (Figure 30, 32).

## Model based indices

Delta-lognormal modeling methods were used to estimate relative abundance indices for greater amberjack (Lo et al. 1992). The main advantage of using this method is allowance for the probability of zero catch (Ortiz et al. 2000). The index computed by this method is a mathematical combination of yearly abundance estimates from two distinct generalized linear models: a binomial (logistic) model which describes proportion of positive abundance values (i.e. presence/absence) and a lognormal model which describes variability in only the nonzero abundance data (Lo et al. 1992).

The delta-lognormal index of relative abundance ( $I_{y}$ ) as described by Lo et al. (1992) was estimated as:
(2) $I_{y}=c_{y} p_{y}$,
where $c_{y}$ is the estimate of mean CPUE for positive catches only for year $y$, and $p_{y}$ is the estimate of mean probability of occurrence during year $y$. Both $c_{y}$ and $p_{y}$ were estimated using generalized linear models. Data used to estimate abundance for positive catches (c) and probability of occurrence $(p)$ were assumed to have a lognormal distribution and a binomial distribution, respectively, and modeled using the following equations:

$$
\begin{equation*}
\ln (c)=X \beta+\varepsilon \tag{3}
\end{equation*}
$$

and

$$
\begin{equation*}
p=\frac{e^{\mathrm{X} \beta+\varepsilon}}{1+e^{\mathrm{X} \beta+\varepsilon}}, \tag{4}
\end{equation*}
$$

respectively, where $c$ is a vector of the positive catch data, $p$ is a vector of the presence/absence data, $X$ is the design matrix for main effects, $\beta$ is the parameter vector for main effects, and $\varepsilon$ is a vector of independent normally distributed errors with expectation zero and variance $\sigma^{2}$.
Therefore, $c_{y}$ and $p_{y}$ were estimated as least-squares means for each year along with their corresponding standard errors, $\operatorname{SE}\left(c_{y}\right)$ and $\operatorname{SE}\left(p_{y}\right)$, respectively. From these estimates, $I_{y}$ was calculated, as in equation (1), and its variance calculated as:
(5) $\quad V\left(I_{y}\right) \approx V\left(c_{y}\right) p_{y}^{2}+c_{y}^{2} V\left(p_{y}\right)+2 c_{y} p_{y} \operatorname{Cov}(c, p)$,
where:
(6) $\left.\quad \operatorname{Cov}(c, p) \approx \rho_{\mathrm{c}, \mathrm{p}} \mid \operatorname{SE}\left(c_{y}\right) \operatorname{SE}\left(p_{y}\right)\right]$,
and $\rho_{\mathrm{c}, \mathrm{p}}$ denotes correlation of $c$ and $p$ among years.
The submodels of the delta-lognormal model were built using a backward selection procedure based on type 3 analyses with an inclusion level of significance of $\alpha=0.05$. Binomial submodel performance was evaluated using AIC, while the performance of the lognormal submodel was evaluated based on analyses of residual scatter and QQ plots in addition to AIC. Variables that could be included in the submodels were: Year (1987-2011).

## Design based indices

A delta-lognormal modeling approach (Lo et al., 1992) was used to develop abundance indices. Independent variables used in the model were year, region and depth. Region is divided into east and west at 89.15 west longitude. The GENMOD procedure in SAS (v.9.2) was used to conduct separate forward stepwise regressions on the binomial and lognormal sub-models to determine which variables to retain for use in fitting the delta lognormal model. Only variables that reduced model deviance by at least $1 \%$ with a type 3 analysis level of significance of $\alpha=0.05$ were retained. The GLIMMIX and MIXED procedures in SAS (v. 9.2) were used to develop the binomial and lognormal sub-models, respectively. A backward selection procedure was used to determine which variables retained from the GENMOD procedure were to be included into each final sub-model based on a type 3 analyses with a level of significance for inclusion of $\alpha=0.01$. Year was including in all terminal models regardless of significance, while region and depth were retained in both the binomial and lognormal sub-models. The estimates from each model were weighted using the stratum area, and separate covariance structures were developed for each survey year. For the binomial models, a logistic-type mixed model was employed.

## Results

Greater amberjack were observed at banks in both the western and eastern GOM (Figures $3-17$ ), and the spatial distributions observed are highly reflective of the reef sampling universe used to select sampling sites (Figures 1-2). Gaps in habitat level information exist in central Florida, Mississippi river delta region, and portions of the Texas coast. In most years the survey shows good coverage in the defined sampling universe, and coverage improved through time as the sampling universe expanded and more sites were added to the survey. Reef blocks from coastal Texas are often not selected for sampling due to small spatial coverage of reef, and frequent high winds and rough sea states during the spring/early summer sampling season.

Design based analysis retained year, depth, and maxrelief in the binomial and log-normal GOM-wide sub-model. Design based greater amberjack proportion positives ranged from 0.15 (1997) to 0.34 (2002) with a reported value of 0.26 in 2012 (Figure 18), while standardized index of abundance ranged from 0.61 (1997) to 1.84 (2002), and reported a value of 0.94 in 2012 (Table 2, Figure 20). Coefficient of variation ranged from $10 \%$ (2012) to $19 \%$ (1993), with the lowest values having been reported in the most recent survey year.

Design based analysis retained year, depth, and maxrelief in the binomial and log-normal east-GOM sub-model. Design based east-GOM greater amberjack proportion positives ranged from 0.08 (1997) to 0.30 (2002) (Figure 26), and the standardized index of abundance ranged from 0.53 (1997) to 2.21 (2002) (Table 5, Figure 28). Coefficient of variation ranged from $18.3 \%$ (2012) to $43 \%$ (1995) with the lowest values having been reported in the most recent survey year.

Design based analysis retained year, region and depth in the binomial and log-normal west-GOM sub-model. Design based greater amberjack proportion positive ranged from 0.12 (2004) to 0.47 (1995) with a reported value of 0.29 in 2012 (Figure 34), while standardized index values ranged from 0.36 (1996) to 1.66 (1995) (Table 8, Figure 36). Coefficient of variation ranged from $14.7 \%$ (2012) to $44.7 \%$ (2004), with the lowest values having been reported in the most recent survey year.

Proportion positives, lo-index, and standardized index values from in 2012 are average relative to all other sample years in the survey. Gulf wide and east GOM values for proportion positives, lo-index, and standardized index output suggest that gulf wide population trends appear to be in large part driven by eastern populations. Western indices appear to be slightly out of sync with several peak years not match the trends than observed in the eastern GOM. Median lengths 'notched' boxplots suggest that greater amberjack sampled in 2012 are significantly larger than any other year sampled, although the range of lengths appears to be within normal bounds observed over the history of the survey.

## Literature cited

Cochran, W.G. 1977. Sampling Techniques. John Wiley \& Sons. New York, NY. 428 p.
Lo, N. C. H., L.D. Jacobson, and J.L. Squire. 1992. Indices of relative abundance from fish spotter data based on delta-lognormal models. Can. J. Fish. Aquat. Sci. 49: 2515-1526.

Figure 3. Spatial distribution of greater amberjack observed and associated min-count values during the 1993 reef fish video survey.


Figure 4. Spatial distribution of greater amberjack observed and associated min-count values during the 1994 reef fish video survey.


Figure 5. Spatial distribution of greater amberjack observed and associated min-count values during the 1995 reef fish video survey.


Figure 6. Spatial distribution of greater amberjack observed and associated min-count values during the 1996 reef fish video survey.


Figure 7. Spatial distribution of greater amberjack observed and associated min-count values during the 1997 reef fish video survey.


Figure 8. Spatial distribution of greater amberjack observed and associated min-count values during the 2002 reef fish video survey.


Figure 9. Spatial distribution of greater amberjack observed and associated min-count values during the 2004 reef fish video survey.


Figure 10. Spatial distribution of greater amberjack observed and associated min-count values during the 2005 reef fish video survey.


Figure 11. Spatial distribution of greater amberjack observed and associated min-count values during the 2006 reef fish video survey.


Figure 12. Spatial distribution of greater amberjack observed and associated min-count values during the 2007 reef fish video survey.


Figure 13. Spatial distribution of greater amberjack observed and associated min-count values during the 2008 reef fish video survey.


Figure 14. Spatial distribution of greater amberjack observed and associated min-count values during the 2009 reef fish video survey.


Figure 15. Spatial distribution of greater amberjack observed and associated min-count values during the 2010 reef fish video survey.


Figure 16. Spatial distribution of greater amberjack observed and associated min-count values during the 2011 reef fish video survey.


Figure 17. Spatial distribution of greater amberjack observed and associated min-count values during the 2012 reef fish video survey.


Table 1. Iteration history (a), fit statistics (b), type III tests (c), and over-dispersion diagnostics (d) of the GLIMMIX binomial on proportion positives for the GOM-wide model.
a

| Iteration History |  |  |  |
| ---: | ---: | ---: | ---: |
| Iteration | Evaluations | -2 Res Log Like | Criterion |
| 1 | 1 | 21641.97029372 | 0.00000000 |

b

| Fit Statistics |  |
| :--- | :--- |
| -2 Res Log Likelihood | 21642.0 |
| AIC (smaller is better) | 21672.0 |
| AICC (smaller is better) | 21672.1 |
| BIC (smaller is better) | 21768.4 |

c $\qquad$
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Effect | $D F$ | $D F$ | Chi-Square | F Value | $\operatorname{Pr}>$ ChiSq | $\operatorname{Pr}>F$ |
| year | 14 | 1409 | 75.18 | 5.34 | $<.0001$ | $<.0001$ |
| depth | 1 | 3713 | 46.82 | 46.82 | $<.0001$ | $<.0001$ |
| MAXRELIEF | 1 | 3233 | 28.36 | 28.36 | $<.0001$ | $<.0001$ |

d

| Description | Value |
| :--- | ---: |
| Deviance | 849.4589 |
| Scaled Deviance | 4859.3083 |
| Pearson Chi-Square | 864.0730 |
| Scaled Pearson Chi-Square | 4942.9082 |
| Extra-Dispersion Scale | 0.1748 |

Figure 18. GOM-wide observed versus proportion positive for design based simulation.
Delta lognormal mincount for Amberjack Diagnostic plots: 1) Obs vs Pred Proport Posit


$$
\text { PLOT } \because \text { obppos } \quad \theta-\Delta \Delta \text { bc_pos }
$$

Figure 19. GOM-wide chi-square residuals of proportion positive design based model.
Delta lognormal mincount for Amberjack Chisq Residuals proportion positive


Table 2. GOM-wide greater amberjack lo and standardized index of abundance values by year design based model.

| SurveyYear | Frequency | $N$ | LoIndex | StdIndex | SE | $C V$ | LCL | UCL |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1993 | 0.15723 | 159 | 0.47200 | 1.14831 | 0.090376 | 0.19147 | 0.78568 | 1.67833 |
| 1994 | 0.27966 | 118 | 0.49831 | 1.21231 | 0.089951 | 0.18051 | 0.84738 | 1.73441 |
| 1995 | 0.29204 | 113 | 0.45749 | 1.11299 | 0.081160 | 0.17740 | 0.78270 | 1.58268 |
| 1996 | 0.15260 | 308 | 0.28655 | 0.69713 | 0.045083 | 0.15733 | 0.50991 | 0.95310 |
| 1997 | 0.14591 | 281 | 0.25085 | 0.61028 | 0.046058 | 0.18361 | 0.42401 | 0.87840 |
| 2002 | 0.34109 | 258 | 0.75454 | 1.83568 | 0.082089 | 0.10879 | 1.47768 | 2.28043 |
| 2004 | 0.18182 | 198 | 0.39662 | 0.96491 | 0.061284 | 0.15452 | 0.70969 | 1.31192 |
| 2005 | 0.22308 | 390 | 0.41865 | 1.01852 | 0.044949 | 0.10736 | 0.82221 | 1.26171 |
| 2006 | 0.14925 | 402 | 0.30351 | 0.73839 | 0.041943 | 0.13819 | 0.56081 | 0.97219 |
| 2007 | 0.17521 | 468 | 0.36762 | 0.89436 | 0.043879 | 0.11936 | 0.70503 | 1.13454 |
| 2008 | 0.16438 | 292 | 0.30484 | 0.74163 | 0.047110 | 0.15454 | 0.54544 | 1.00839 |
| 2009 | 0.22087 | 412 | 0.44078 | 1.07234 | 0.048554 | 0.11016 | 0.86088 | 1.33575 |
| 2010 | 0.23549 | 293 | 0.34333 | 0.83526 | 0.044116 | 0.12850 | 0.64665 | 1.07888 |
| 2011 | 0.24769 | 432 | 0.48580 | 1.18189 | 0.055936 | 0.11514 | 0.93950 | 1.48682 |
| 2012 | 0.25607 | 453 | 0.38472 | 0.93597 | 0.038541 | 0.10018 | 0.76642 | 1.14304 |

Table 3. Fit statistics (a), and type III tests (b) of the GLM on positive catches for the GOM-wide design based model.
a

| Fit Statistics |  |
| :--- | :--- |
| -2 Res Log Likelihood | 2523.4 |
| AIC (smaller is better) | 2525.4 |
| AICC (smaller is better) | 2525.4 |
| BIC (smaller is better) | 2530.2 |

$\qquad$
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Effect | $D F$ | DF | F Value | $\operatorname{Pr}>F$ |
| year | 14 | 918 | 1.63 | 0.0664 |
| depth | 1 | 918 | 8.14 | 0.0044 |
| MAXRELIEF | 1 | 918 | 0.01 | 0.9255 |

Figure 20. GOM-wide observed versus standardized mincount for design based model.
Delta lognormal mincount for Amberjack
Observed and Standardized mincount (95\% CI)


Figure 21. GOM-wide observed versus predicted mincount of positive data for design based model.

Delta lognormal mincount for Amberjack
Diagnostic plots: 2) Obs vs Pred mincount of Posit only MINCOUNT


Figure 22. GOM-wide observed versus predicted mincount for design based model.
Delta lognormal mincount for Amberjack Diagnostic plots: 3) Obs vs Pred mincount Input units


PLOT $\because M I N C O U N T \quad \Leftrightarrow-\otimes \otimes$ index

Figure 23. GOM wide residuals of positive mincounts by year for design based model.
Delta lognormal mincount for Amberjack Residuals positive mincounts * Year


Figure 24 GOM-wide residuals distribution from positive mincount design based model.
Delta lognormal mincount for Amberjack
Residuals positive mincount Distribution


Figure 25 GOM-wide qqplot of residuals of positive mincounts from design based model.
Delta lognormal mincount for Amberjack QQplot Residuals Positive mincount rates


Table 4. Iteration history (a), fit statistics (b), type III tests (c), and over-dispersion diagnostics (d) of the GLIMMIX binomial on proportion positives for the east GOM model.
a

| Iteration History |  |  |  |
| ---: | ---: | ---: | ---: |
| Iteration | Evaluations | -2 Res Log Like | Criterion |
| 1 | 1 | 13406.28031850 | 0.00000000 |

b

| Fit Statistics |  |
| :--- | :--- |
| -2 Res Log Likelihood | 13406.3 |
| AIC (smaller is better) | 13436.3 |
| AICC (smaller is better) | 13436.5 |
| BIC (smaller is better) | 13525.6 |

c $\qquad$
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Effect | $D F$ | DF | Chi-Square | F Value | $\operatorname{Pr}>$ ChiSq | $\operatorname{Pr}>F$ |
| year | 14 | 857 | 45.95 | 3.25 | $<.0001$ | $<.0001$ |
| depth | 1 | 2183 | 54.50 | 54.50 | $<.0001$ | $<.0001$ |
| MAXRELIEF | 1 | 1951 | 39.53 | 39.53 | $<.0001$ | $<.0001$ |

d

| Description | Value |
| :--- | ---: |
| Deviance | 846.8134 |
| Scaled Deviance | 2884.3447 |
| Pearson Chi-Square | 911.3169 |
| Scaled Pearson Chi-Square | 3104.0509 |
| Extra-Dispersion Scale | 0.2936 |

Figure 26. Observed versus predicted proportion positive from east GOM design based model.
Delta lognormal mincount for Amberjack
Diagnostic plots: 1) Obs vs Pred Proport Posit


PLOT $\because$ obppos $\forall-\otimes$ bc_pos
Figure 27. Chi-square residuals of proportion positives from east GOM design based model.
Delta lognormal mincount for Amberjack
Chisq Residuals proportion positive


Table 5. East GOM greater amberjack lo and standardized index of abundance by year for design based model.

| SurveyYear | Frequency | $N$ | Lolndex | StdIndex | SE | $C V$ | $L C L$ | $U C L$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1993 | 0.14035 | 114 | 0.64732 | 1.53622 | 0.20248 | 0.31280 | 0.83382 | 2.83033 |
| 1994 | 0.20000 | 75 | 0.45663 | 1.08369 | 0.16067 | 0.35185 | 0.54722 | 2.14606 |
| 1995 | 0.14516 | 62 | 0.26972 | 0.64011 | 0.11642 | 0.43163 | 0.28005 | 1.46308 |
| 1996 | 0.16058 | 137 | 0.38681 | 0.91799 | 0.10971 | 0.28361 | 0.52631 | 1.60114 |
| 1997 | 0.08108 | 148 | 0.22434 | 0.53242 | 0.08965 | 0.39960 | 0.24657 | 1.14965 |
| 2002 | 0.29814 | 161 | 0.93279 | 2.21372 | 0.18087 | 0.19390 | 1.50750 | 3.25078 |
| 2004 | 0.20134 | 149 | 0.41616 | 0.98763 | 0.09836 | 0.23636 | 0.61956 | 1.57437 |
| 2005 | 0.20472 | 254 | 0.40330 | 0.95712 | 0.07432 | 0.18428 | 0.66410 | 1.37941 |
| 2006 | 0.11278 | 266 | 0.23918 | 0.56761 | 0.05935 | 0.24814 | 0.34812 | 0.92548 |
| 2007 | 0.16452 | 310 | 0.33847 | 0.80325 | 0.07119 | 0.21033 | 0.52983 | 1.21778 |
| 2008 | 0.12426 | 169 | 0.30125 | 0.71493 | 0.08757 | 0.29068 | 0.40447 | 1.26369 |
| 2009 | 0.20325 | 246 | 0.49640 | 1.17806 | 0.09672 | 0.19485 | 0.80076 | 1.73314 |
| 2010 | 0.21939 | 196 | 0.35834 | 0.85042 | 0.07642 | 0.21325 | 0.55779 | 1.29659 |
| 2011 | 0.23585 | 318 | 0.45902 | 1.08936 | 0.09148 | 0.19930 | 0.73409 | 1.61657 |
| 2012 | 0.22925 | 253 | 0.39082 | 0.92749 | 0.07153 | 0.18303 | 0.64512 | 1.33345 |

Table 6. Fit statistics (a), and type III tests (b) of the GLM on positive catches for the east GOM design based model.
a

| Fit Statistics |  |
| :--- | ---: |
| -2 Res Log Likelihood | 1372.4 |
| AIC (smaller is better) | 1374.4 |
| AICC (smaller is better) | 1374.4 |
| BIC (smaller is better) | 1378.6 |

b
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Effect | DF | DF | F Value | $\operatorname{Pr}>F$ |
| year | 14 | 489 | 1.17 | 0.2913 |
| depth | 1 | 489 | 3.35 | 0.0680 |
| MAXRELIEF | 1 | 489 | 0.36 | 0.5475 |

Figure 28. Observed and standardized mincounts from east GOM design based model.
Delta lognormal mincount for Amberjack Observed and Standardized mincount (95\% Cl)

$\begin{array}{lll}\text { PLOT } & \leftrightarrow \text { STDmincount } & \forall \diamond \text { LCI } \\ & \forall \Leftarrow \text { UCI } & \bullet \text { obsmincount }\end{array}$

Figure 29. Observed versus predicted mincounts from east GOM design based model.
Delta lognormal mincount for Amberjack
Diagnostic plots: 2) Obs vs Pred mincount of Posit only MINCOUNT


Figure 30. Observed versus predicted mincounts from east GOM design based models
Delta lognormal mincount for Amberjack Diagnostic plots: 3) Obs vs Pred mincount Input units


PLOT $\because$ MINCOUNT $\Leftrightarrow-\otimes \otimes$ index

Figure 31 Residuals of positive mincounts for east GOM design based model.
Delta lognormal mincount for Amberjack
Residuals positive mincounts * Year


Figure 32. Positive mincount distribution from east GOM design based model.
Delta lognormal mincount for Amberjack
Residuals positive mincount Distribution


Figure 33. QQ plot of positive mincounts from east GOM design based model.
Delta lognormal mincount for Amberjack QQplot Residuals Positive mincount rates


Table 7. Iteration history (a), fit statistics (b), type III tests (c), and over-dispersion diagnostics (d) of the GLIMMIX binomial on proportion positives for the west GOM model.
a

| Iteration History |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
| Iteration | Evaluations | -2 Res Log Like | Criterion |
| 1 | 1 | 8161.95087073 | 0.00000000 |

b

| Fit Statistics |  |
| :--- | :--- |
| -2 Res Log Likelihood | 8162.0 |
| AIC (smaller is better) | 8192.0 |
| AICC (smaller is better) | 8192.2 |
| BIC (smaller is better) | 8273.7 |

c $\qquad$
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Effect | $D F$ | $D F$ | Chi-Square | F Value | $\operatorname{Pr}>$ ChiSq | $\operatorname{Pr}>F$ |
| year | 14 | 490 | 54.97 | 3.86 | $<.0001$ | $<.0001$ |
| depth | 1 | 1357 | 8.27 | 8.27 | 0.0040 | 0.0041 |
| MAXRELIEF | 1 | 997 | 0.19 | 0.19 | 0.6588 | 0.6589 |

d

| Description | Value |
| :--- | ---: |
| Deviance | 799.6733 |
| Scaled Deviance | 1899.7182 |
| Pearson Chi-Square | 751.0466 |
| Scaled Pearson Chi-Square | 1784.1997 |
| Extra-Dispersion Scale | 0.4209 |

Figure 34. Observed versus predicted proportion positive from west GOM design based model.
Delta lognormal mincount for Amberjack Diagnostic plots: 1) Obs vs Pred Proport Posit


PLOT $\because$ obppos $\forall-\otimes \diamond$ bc_pos

Figure 35. Chi-square residuals of proportion positives of west GOM design based model.
Delta lognormal mincount for Amberjack Chisq Residuals proportion positive


Table 8. West GOM greater amberjack Lo and standardized index of abundance by year for design based model.

| SurveyYear | Frequency | $N$ | Lolndex | StdIndex | $S E$ | $C V$ | $L C L$ | $U C L$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 1993 | 0.20000 | 45 | 0.19660 | 0.44445 | 0.07140 | 0.36315 | 0.21984 | 0.89854 |
| 1994 | 0.41860 | 43 | 0.61393 | 1.38790 | 0.16047 | 0.26139 | 0.82996 | 2.32091 |
| 1995 | 0.47059 | 51 | 0.73571 | 1.66321 | 0.15628 | 0.21242 | 1.09264 | 2.53174 |
| 1996 | 0.14620 | 171 | 0.16059 | 0.36303 | 0.04264 | 0.26555 | 0.21538 | 0.61190 |
| 1997 | 0.21805 | 133 | 0.29453 | 0.66583 | 0.07043 | 0.23913 | 0.41547 | 1.06705 |
| 2002 | 0.41237 | 97 | 0.53543 | 1.21043 | 0.09404 | 0.17564 | 0.85416 | 1.71531 |
| 2004 | 0.12245 | 49 | 0.26854 | 0.60708 | 0.12007 | 0.44714 | 0.25848 | 1.42584 |
| 2005 | 0.25735 | 136 | 0.52155 | 1.17906 | 0.09489 | 0.18195 | 0.82183 | 1.69156 |
| 2006 | 0.22059 | 136 | 0.63467 | 1.43479 | 0.12416 | 0.19563 | 0.97378 | 2.11405 |
| 2007 | 0.19620 | 158 | 0.51631 | 1.16721 | 0.10005 | 0.19378 | 0.79503 | 1.71363 |
| 2008 | 0.21951 | 123 | 0.37467 | 0.84701 | 0.08165 | 0.21793 | 0.55054 | 1.30312 |
| 2009 | 0.24699 | 166 | 0.38130 | 0.86199 | 0.06781 | 0.17784 | 0.60567 | 1.22679 |
| 2010 | 0.26804 | 97 | 0.29601 | 0.66919 | 0.07528 | 0.25433 | 0.40560 | 1.10409 |
| 2011 | 0.28070 | 114 | 0.66394 | 1.50096 | 0.12004 | 0.18080 | 1.04855 | 2.14855 |
| 2012 | 0.29000 | 200 | 0.44139 | 0.99785 | 0.06506 | 0.14741 | 0.74425 | 1.33787 |

Table 9. Fit statistics (a), and type III tests (b) of the GLM on positive catches for the west GOM design based model.
a

| Fit Statistics |  |
| :--- | ---: |
| -2 Res Log Likelihood | 946.1 |
| AIC (smaller is better) | 948.1 |
| AICC (smaller is better) | 948.1 |
| BIC (smaller is better) | 952.1 |

$\qquad$
Type 3 Tests of Fixed Effects

|  | Num | Den |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Effect | $D F$ | $D F$ | $F$ Value | $\operatorname{Pr}>F$ |
| year | 14 | 411 | 2.55 | 0.0016 |
| depth | 1 | 411 | 0.02 | 0.8849 |
| MAXRELIEF | 1 | 411 | 1.62 | 0.2042 |

Figure 36. Observed and standardized mincounts from west GOM design based model.
Delta lognormal mincount for Amberjack Observed and Standardized mincount (95\% Cl)

## STDmincount



PLOT $\leftrightarrow$ STDmincount $\forall \otimes \forall \mathrm{LCl}$
$\Leftrightarrow \Leftrightarrow \mathrm{UCl} \quad \bullet$ obsmincount
Figure 37. Observed versus predicted mincounts of positive data from west GOM design based model.

Delta lognormal mincount for Amberjack
Diagnostic plots: 2) Obs vs Pred mincount of Posit only MINCOUNT


Figure 38. Observed versus predicted mincounts of west GOM design based model.
Delta lognormal mincount for Amberjack
Diagnostic plots: 3) Obs vs Pred mincount Input units


PLOT $\because M I N C O U N T ~ \&-\Delta \Delta$ index

Figure 39. Residuals of positive mincounts for west GOM design based model.
Delta lognormal mincount for Amberjack
Residuals positive mincounts * Year


Figure 40. Positive mincount distribution of residuals for west GOM design based model.
Delta lognormal mincount for Amberjack Residuals positive mincount Distribution


Figure 41. QQ plot of positive mincounts for west GOM design based model.
Delta lognormal mincount for Amberjack QQplot Residuals Positive mincount rates


Figure 42. Greater amberjack length frequency of fish measured from video with lasers in 1995.


Figure 43. Greater amberjack length frequency of fish measured from video with lasers in 1996.


Figure 44. Greater amberjack length frequency of fish measured from video with lasers in 1997.


Figure 45. Greater amberjack length frequency of fish measured from video with lasers in 2001.


Figure 46. Greater amberjack length frequency of fish measured from video with lasers in 2002.


Figure 47. Greater amberjack length frequency of fish measured from video with lasers in 2004.


Figure 48. Greater amberjack length frequency of fish measured from video with lasers in 2005.


Figure 49. Greater amberjack length frequency of fish measured from video with lasers in 2006.


Figure 50. Greater amberjack length frequency of fish measured from video with lasers in 2007.


Figure 51. Greater amberjack length frequency of fish measured with stereo cameras in 2008.


Figure 52. Greater amberjack length frequency of fish measured with stereo cameras in 2009.


Figure 53. Greater amberjack length frequency of fish measured with stereo cameras in 2010.


Figure 54. Greater amberjack length frequency of fish measured with stereo cameras in 2011.


Figure 55. Greater amberjack length frequency of fish measured with stereo cameras in 2012.


Figure 56. Greater amberjack mean lengths by year. Upper and lower quartiles represented within boxes, whiskers extend to subsequent quartiles, and non-overlapping notches indicate groups for which median responses are likely different.


